

Short title: CHILDHOOD ABUSE AMONG HOMELESS PEOPLE

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**Prevalence of childhood abuse among people who are homeless in Western countries: A systematic review and meta-analysis**

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Abstract

Purpose: This article systematically reviews studies of prevalence of childhood experience of physical and sexual abuse in adult people who are homeless in Western countries.

Methods: Medline, PsychInfo, and the Cochrane Library were searched using the keywords: homeless\*, child\* abuse, child\* trauma, and child\* adversity and the bibliographies of identified articles were reviewed. Sources of heterogeneity in the prevalence rates were explored by meta-regression analysis

Results: Twenty-four reports published between January 1990 and August 2013 in three countries provided estimates obtained from up to 9,730 adult individuals who were homeless.

Prevalence of reported childhood physical abuse ranged from 6 to 94% with average prevalence of 37%, 95% CI [25, 51]. Reported sexual abuse ranged from 4 to 62%, with average prevalence estimated as 32%, 95% CI [23, 44] for females and 10% for males, 95% CI [6, 17]. Substantial heterogeneity was observed among the studies ( $I^2 \geq 98\%$ ). Including moderators greatly reduced but did not eliminate this heterogeneity. Moderator analyses suggested that reported physical abuse tended to be higher for predominately white samples and tended to be lower for younger samples. Sexual abuse was far more prevalent in predominately female samples and slightly higher in non-US samples and convenience samples.

Conclusions: The findings of this study suggest that childhood physical and sexual abuse is more prevalent among the homeless in Western countries than in the global population. Physical abuse appears to be particularly prevalent in younger samples and sexual abuse rates are higher in predominately female samples. Further investigation is needed to advance our understanding of how trauma informed treatment and care for the homeless effectively can take into account the service user's experiences of childhood abuse.

**Keywords:** childhood physical abuse; childhood sexual abuse; adult people who are  
homeless; meta regression analysis

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**Prevalence of childhood abuse among people who are homeless in Western countries: A systematic review and meta-analysis.**

INTRODUCTION

Over the last forty years, it has been increasingly recognised that homelessness is a multifaceted problem. Unfavourable structural conditions, for example a shortage of affordable housing and unemployment, have been identified as prerequisites that explain widespread homelessness (broadly defined to include rough sleeping, living in emergency or insecure accommodation) [1-3]. Furthermore, because the ability to access employment is limited by the lack of a stable address, homelessness is associated with poverty [4,5]. Third, people who are homeless often experience severe difficulties with housing related services, health services and education [6,7]. For many, these problems are sustained or exacerbated by poor physical and mental health and the need to support alcohol or drug dependencies [2,8]. In fact, a recent review article investigated the mortality rate among the homeless in Boston and found that the all-cause mortality rate remains high and unchanged since 1988 to 1993 [9].

Despite the expansion of research on issues that cause and surround homelessness, the number of people experiencing homelessness in most, if not all, Western countries has increased or remained stable during the past decade [10-12]. There is also an increase in the number of people with mental illnesses living on the streets or in shelters or hostels [13,14] and a nationwide, prospective, register-based cohort study of homeless people in Denmark showed that registered substance abuse disorder was associated with the highest mortality risk compared with no psychiatric contact registered [15]. Several review articles have confirmed that psychiatric disorders are more prevalent among people who are homeless in

1 Western countries than in the age-matched general population and they are more likely to  
2 have alcohol and drug dependence [16-20].  
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4 The role of traumatic experiences in the homeless is a pertinent issue that has been studied  
5 by many researchers. Some authors have used Seligman's theory of learned helplessness [21]  
6 to explain how, for many, becoming homeless is a traumatic event in itself (22-24). This is  
7 because, "like other traumas, becoming homeless frequently renders people unable to control  
8 their daily lives" [23, p 122] In addition and parallel with the body of research demonstrating  
9 that experiences of childhood abuse in members of the general population is linked to, e.g.,  
10 mental health problems and drug and alcohol abuse later in life [25-26], many researchers  
11 have investigated the role of child maltreatment in adult people who are homeless. Four types  
12 of child maltreatment are generally recognised: physical abuse; sexual abuse; emotional  
13 abuse; and neglect. However, the impact of childhood physical abuse on health and wellbeing  
14 later in life has been studied less than childhood sexual abuse, and childhood emotional abuse  
15 and neglect has received the least scientific attention [26-30].  
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33 Several reasons for this "neglect of neglect" [31, p 530] has been discussed in the  
34 literature, most importantly a lack of consistency and clarity regarding definitions of neglect  
35 [27,29,32] and the overlap between sexual and physical childhood abuse and neglect. For  
36 example, one study reported that almost half of the neglected children were also victims of  
37 physical abuse and about 21% were also sexually abused [33]. To our knowledge, there have  
38 been no prior reviews of experiences of childhood sexual and childhood physical abuse in  
39 adult individuals who are homeless. This systematic review aims to provide a comprehensive  
40 description of the current knowledge of homeless adults and childhood experiences of  
41 physical and sexual abuse. Further, disparities across gender, age, and geographic regions are  
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1 examined. Directions for future research on childhood trauma and homelessness are also  
2 identified.  
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## 7 **METHOD**

### 8 *Selection of studies*

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11 Using Psychinfo, MEDLINE, the Cochrane Library and a direct library search we sought  
12 relevant articles published between January 1990 and August 2013. The search terms that  
13 were used include homeless\*, child\* abuse, child\* trauma, and child\* adversity. The  
14 references cited in the articles that are included in this review were also manually examined  
15 to further identify any additional relevant studies. A flow chart summarizing the selection  
16 process is presented in Figure 1.  
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27 The inclusion criteria required that the studies: i) have a naturalistic research design with  
28 data collection predominately from adult individuals who were currently experiencing, or  
29 who had previously experienced, homelessness, ii) be published between January 1990 -  
30 August 2013 in English language, peer-reviewed journals, and iii) include separate reports of  
31 the prevalence of childhood sexual abuse (CSA), childhood physical abuse (CPA) or both  
32 CSA and CPA in people who were, or previously had been, homeless. Samples were  
33 considered to predominately adult if all individuals were aged 16 or over and if the mean age  
34 was reported or estimated as 18 or greater. A protocol was used to systematically extract  
35 information from each eligible study.<sup>1</sup>  
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49 As can be seen in Figure 1, a total of 98 studies were identified from database searches  
50 and a manual search. Of these, 48 studies were deemed not relevant (27 studies did not report  
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56 <sup>1</sup> The protocol is available from:

57 [http://www2.ntpsychology.net/~baguley/Sundin\\_Baguley\\_protocol\\_160813.doc](http://www2.ntpsychology.net/~baguley/Sundin_Baguley_protocol_160813.doc)  
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1 CPA and/or CSA data, and 21 studies reported on children or adolescents). In addition,  
2 studies were excluded that only identified a combined abuse score (physical and sexual  
3 abuse) and those who only included a select sample of homeless people (mothers or pregnant  
4 women; veterans; homeless people who were referred to a mental health intervention). A total  
5 of 26 studies were retained. However, detailed examination of the data (including sample  
6 characteristics and the dates and location of data collection) indicated that two pairs of studies  
7 had samples that may have overlapped. In each case the study that reported a larger data set  
8 was included. Consequently, the study by Davis and Winkleby [34] was included in  
9 preference to Winkleby and Fleshin [35], and Stein et al. [36] was included in preference to  
10 Nyamathi et al. [37].  
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31 All variables of the remaining 24 studies that are relevant for this study are summarized in  
32 Table 1 [34,36, 38-59] and, where sufficient information is available, percentage CSA and  
33 CPA data are presented for distinct samples (e.g., males and females; whites and non-whites)  
34 within each study.  
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48 *Study characteristics*  
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50 For studies where both CPA and CSA data are reported,  $n$  is usually identical for each  
51 measure, but in a few instances sample sizes differ for CPA and CSA (e.g., presumably  
52 because of infrequent events such as non-disclosure or recording error). For this reason  $n$  is  
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1 reported separately for CSA and CPA in Table 1.<sup>2</sup> For each measure the percentage  
2 prevalence rate is also reported. Ten studies used randomization ( $n = 3,768$ ), twelve studies  
3 had convenience samples ( $n = 3,118$ ) and two employed some form of stratification ( $n =$   
4  $2,373$ ). In total 9,730 homeless individuals were included in the sample, a majority of those  
5 were recruited from shelters or hostels for the homeless and some were recruited from day  
6 centres, soup kitchens, missions or the street.  
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14 The studies included in the sample were published between 1990 and 2013; of those,  
15 10 studies were published before 2000 and 15 since 2000 (though the earliest reported date of  
16 data collection is 1985). Most of the data are North American in origin; 18 studies collected  
17 data in the US ( $n = 8,230$ ), 2 in Canada ( $n = 814$ ), and 2 in the UK ( $n = 215$ ). Figure 2 shows  
18 the distribution of sample sizes ( $n_{CSA}$ ) across the 24 studies. Most are relatively small (with  $n$   
19  $< 200$ ) and only two studies have total  $n$  greater than 1,000 [34,48]. The majority of the  
20 studies were of urban samples, although three took national (US) samples and six also  
21 included non-urban (rural or suburban) locations. There was also considerable variability  
22 between studies in measures used to determine prevalence. Most involved a single interview  
23 or questionnaire item, however, five studies use psychometric instruments [42-43,46,52,57].  
24 This included four using the Childhood Trauma Questionnaire (CTQ) [42-43,52,57]. Of these  
25 one used the low to moderate cut-off [42] as an indicator for reported abuse, two used the  
26 moderate to extreme cut-off [43,57], while the fourth study reported prevalence data for both  
27 cut-offs (and was coded for meta-analysis according to the stricter criterion) [52]. Study  
28 quality and biases arising from poor quality studies were assessed by inclusion of potential  
29 moderators notably those relating to type of sample (e.g., convenience versus random or  
30 stratified) and year of study. In addition, to assess the influence variation in the method of  
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56 <sup>2</sup> Unless otherwise stated subsequent reference in the text is to  $n_{CSA}$  (which is generally the larger of  
57 the two sample sizes). All statistical models use  $n_{CPA}$  or  $n_{CSA}$  as appropriate.  
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1 obtaining prevalence estimates criteria for reporting abuse were coded either as lenient  
2 (including using a low to moderate cut-off for the CTQ) or strict (including a moderate to  
3 extreme cut-off for the CTQ) or neither (including where precise wording of the  
4 questionnaire or interview question was not reported). A lenient criterion was also coded if  
5 only one of multiple items relating to different events were endorsed or if a very broad  
6 definition of abuse was used (e.g., “any childhood physical violence” for CPA).  
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22 Five studies provide separate CPA or CSA prevalence (or both) for distinct samples (e.g.,  
23 male and female or white and non-white). Davis and Winkleby [34] report both CPA and  
24 CSA separately for white and non-white samples. Gwadz et al. [43] and North et al. [51] each  
25 report both CPA and CSA prevalence rates for males and females. Johnson et al. [45] also  
26 provide information on distinct male and female samples, but for CSA only. Koegel et al.  
27 [48] report CPA prevalence for the total sample, but provide separate male and female reports  
28 of CSA prevalence.  
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38 The proportion of homeless males across all samples weighted by sample size was .63 and  
39 the weighted proportion of non-white individuals was .61.<sup>3</sup> Four studies, D’Ercole and  
40 Struening [40], Rayburn et al. [53], Stein et al. [36], and Zlotnick et al. [59] sampled only  
41 female participants ( $n = 1,708$ ), while three, Johnson et al. [45], Kim et al. [47], and Sumerlin  
42 [55], had all male samples ( $n = 599$ ). The weighted mean age across all samples was  
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53 <sup>3</sup> The non-white category included any sample or subsample identified as not of white European  
54 origin (e.g., Afro-Caribbean, Asian, and Latino or Hispanic as well as Aborigine and Native American  
55 individuals). In most cases this information was taken from published reports (including other studies  
56 using the same data), but for a few samples was taken from relevant census data for the location and  
57 year of data collection.  
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estimated as 35.9 years ( $SD = 9.7$  years).<sup>4</sup> For males the mean age was estimated as 37.1 years ( $SD = 10.2$ ), while for females it was 33.8 years ( $SD = 8.9$ ).

## RESULTS

All meta-analyses were conducted via restricted maximum likelihood estimation (REML) estimation using the *metafor* package [60] in R [61] with restricted maximum likelihood estimation (REML) used to obtain parameter estimates. A number of effect size metrics can be adopted for proportions (including prevalence rates). Meta-analysis of the raw proportions is problematic if prevalence rates range widely – particularly if any are close to 0% or 100% (as is the case here). Standard meta-analytic methods assume a normal or near-normal distribution of sampling error and constant error variance. The constant variance assumption is likely to be severely violated for prevalence rates where the variance is a function of the rate (being higher for proportions close to the upper or lower limit). For this reason the preferred approach is to transform the proportion prior to analysis (e.g., using the logistic transformation or Freeman-Tukey transformation). Here the logistic transformation performs marginally better (resulting in less heterogeneity) as well as having a more intuitive interpretation (in terms of the log odds of sexual or physical abuse). For all analyses a random effects model is preferred (it being implausible that the true proportion is fixed between study locations or over time). Standard meta-analytic approaches assume that the variances of the original samples are known and produce standard errors that are too small. For this reason a correction to the standard errors that involves a change in test statistic from  $z$  to  $t$  was employed [60,62]. Models with moderator effects were compared using Akaike

information criterion (AIC) using maximum likelihood estimation. This measure assesses the

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<sup>4</sup> Age data are taken from published reports or estimated from available information (e.g., the median or the proportion of the sample falling into different age categories).

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2 informativeness of a model [e.g., see 63], and advantages simpler models (i.e., those with  
3 fewer parameters).  
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### 6 7 *Meta-analysis of CPA* 8

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10 The initial analysis included no moderator variables and included all 26 samples that reported  
11 the proportion of CPA. In this analysis the weighted mean estimate of the log odds of CPA  
12 was -0.520, 95% CI [-1.099, 0.058], corresponding to a proportion of 37%, 95% CI [25, 51],  
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14 AIC = 95.5. Figure 3a shows a forest plot of the proportions (on an untransformed scale),  
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17 while Figure 3b shows the corresponding funnel plot.  
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29 Both plots suggest substantial variability in the proportion of CPA between samples. Indeed,  
30 the level of heterogeneity in the estimates is very large and  $I^2 = .99$ ,  $Q(25) = 1601.4$ ,  $p <$   
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38 .0001, indicating that almost all the variability in the estimates is attributable to variation  
39 between samples rather than sampling error.

40 Given the diverse composition of the samples and the likelihood that CPA is related to  
41 one or more of these factors, heterogeneity of the estimate is to be expected. It is therefore  
42 instructive to account for some or all of the heterogeneity by modelling the prevalence of  
43 CPA as a function of potential moderator variables – notably demographics. A preliminary  
44 analysis included the year of the study, the proportion of white participants, the proportion of  
45 male participants and average age as predictors. Age and year of study were centred prior to  
46 entry in the model. The resulting model is much more informative than the model with no  
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65 moderators (AIC = 76.8,  $\Delta_{AIC} = -18.6$ ). A reduction in AIC of this magnitude is considered  
very strong evidence in favour of the more informative model. The estimate of residual

1 heterogeneity was also reduced, from  $\tau^2 = 1.95$  to  $\tau^2 = 0.74$ . Thus these four moderators  
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3 account for over sixty percent of between-study heterogeneity ( $R^2_{meta} = .62$ ), although  
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5 considerable heterogeneity remains [see 64]. Three further dichotomous moderators  
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7 reflecting other available study characteristics were added. These characteristics were  
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9 whether the sample was US only, whether it included non-urban areas and whether or not a  
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11 convenience sample was used (rather than stratified or random). Adding these moderators did  
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13 not reduce heterogeneity and resulted in a less informative model (AIC = 81.5,  $\Delta_{AIC} = 4.7$ ).  
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17 For the final set of analyses the presence of interaction effects between the year of  
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19 study, proportion of white, proportion of male and average age moderator effects was  
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21 considered. Adding two-way interactions between all four effects did not improve the model  
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23 relative to the main effects only model (AIC = 78.5,  $\Delta_{AIC} = 1.7$ ).  
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27 A final check was to add predictors coding lenient or strict criteria for CPA. Adding  
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29 the strict criterion, but not the lenient criterion (AIC = 78.0,  $\Delta_{AIC} = 1.1$ ), improved model fit  
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31 (AIC = 72.6,  $\Delta_{AIC} = 4.3$ ) and further reduced heterogeneity ( $R^2_{meta} = .68$ ). The final model  
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33 that we report is therefore the initial model including year of study, key demographic  
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35 variables and the strict criterion for determining CPA. Parameter estimates and confidence  
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37 intervals for this model on the transformed (log odds) scale are presented in Table 2.  
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50 Table 2 indicates that more recent studies show higher prevalence rates (perhaps indicating  
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52 changes in willingness to report CPA over time or changes in how CPA is perceived or  
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54 defined by researchers). Age is also associated with prevalence – with younger samples  
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56 tending to report higher rates of CPA. Although age may be confounded with length of  
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homelessness or frequency of homeless episodes, it seems plausible that CPA is a trigger for homelessness in young adults but plays a more indirect role among older adults. Prevalence of CPA also tends to be higher in predominately white homeless samples, but there is little indication of a difference between males and females. This pattern is illustrated in Figure 4, which illustrates the predicted prevalence of CPA for a 2010 study as a function of the proportion of white participants for all male or all female samples when average age is 18, 35 or 55. Figure 4 plots estimates when the strict criterion for determining CPA is not used. Employing a strict criterion for CPA (using a moderate to severe cut-off on the CTQ) decreased prevalence estimates. For purposes of comparison it is interesting to revise our estimate of prevalence from the overall model assuming that the strict criterion was used in each study (but not adjusting for other covariates). This reduces the prevalence estimate from 37% to 29%, 95% CI [8, 65].

INSERT FIGURE 4 ABOUT HERE

### *Meta-analysis of CSA*

For the 29 samples that report CSA, a meta-analysis with no moderators gives the weighted mean estimate of the log odds of CPA as -1.106, 95% CI [-1.523, -0.688]. AIC for the model is 94.7 and the estimate corresponds to CSA prevalence of .25, 95% CI [.18, .33]. There is also considerable heterogeneity in the prevalence of CSA between samples,  $I^2 = .98$ ,  $Q(28) = 1227.5$ ,  $p < .0001$ . Around 98% of the variation is attributable to between-study differences.

Figure 5 shows a forest and funnel plot for the initial model.

INSERT FIGURE 5 ABOUT HERE

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2 As for CPA, year of study, average age (both centered) and the proportion of white or male  
3 participants were added to the model as moderators. This model was more informative than  
4 the model with no moderators (AIC = 66.3,  $\Delta_{AIC} = -24.5$ ) and also reduced the residual  
5 heterogeneity of the model from  $\tau^2 = 1.168$  to  $\tau^2 = 0.429$ . Adding these moderators thus  
6  
7 accounts for over 60% of between-study variability ( $R^2_{meta} = .63$ ). Adding additional study  
8 characteristics (whether the sample was US only, whether it included non-urban areas and  
9 whether convenience sampling was employed) produced a slightly more informative model  
10 (AIC = 63.9,  $\Delta_{AIC} = -2.3$ ) and accounting for nearly 70% of between-study variability ( $R^2_{meta}$   
11 = .69). According to this model CSA prevalence may be slightly lower for US samples and  
12 slightly higher in convenience samples (relative to stratified or random samples), but there  
13 was no indication that samples including non-urban areas differed from urban samples.  
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17 An interaction model containing two-way interactions between key demographic  
18 factors (average age and the proportions of white and male homeless) was also tested, there  
19 being insufficient degrees of freedom to estimate interactions between all predictors. This  
20 model was less informative than the preceding model (AIC = 69.6,  $\Delta_{AIC} = 5.7$ ).  
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24 The final step in the analysis was to add predictors for lenient and strict CSA criteria.  
25 Neither lenient (AIC = 65.2,  $\Delta_{AIC} = 1.2$ ), nor strict (AIC = 65.9,  $\Delta_{AIC} = 2.0$ ) criteria improved  
26 model fit. The absence of effect of an effect of the strict criterion may reflect the sensitivity  
27 of CSA measurement items to other factors (e.g., rapport with the interviewer or differences  
28 in interpretation of terms such as ‘abuse’). Accordingly, the final preferred model includes  
29 the same predictors as for CPA with the addition of indicators for convenience sampling,  
30 samples of US origin and samples that included non-urban areas, but excluding criterion  
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strictness. Parameter estimates and confidence intervals for this model on the transformed scale are presented in Table 3.

INSERT TABLE 3 ABOUT HERE

The results summarized in Table 3 indicate that recent studies are associated with reports of higher CSA prevalence, matching the pattern found for CPA. The effects of the proportion of white and male participants is, however, different than for CPA. Prevalence is strongly associated with the proportion of females in the sample, but unlike the tendency observed for physical abuse, there was no association with the proportion of white homeless in the sample and also little indication that CSA changed with age. US samples tend to have slightly lower prevalence rates than non-US samples and convenience samples have slightly higher estimates of prevalence (the latter possibly indicating a slight sampling bias).

The most striking finding is that prevalence of CSA is far higher for female homeless, 32.3% 95% CI [22.7, 43.7] than male homeless participants, 10.3% 95% [6.1, 17.1].<sup>5</sup> Figure 6 shows the predicted prevalence of CSA for a 2010 study as a function of the proportion of white homeless by sex and average age. Comparison with Figure 4 reveals the large difference in predicted prevalence between males and females for CSA (in comparison to the modest difference for CPA).

INSERT FIGURE 6 ABOUT HERE

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<sup>5</sup> These estimates were obtained from the final model with all other covariates except the indicator for convenience sampling centered. Thus the estimates reflect the average estimate for a random or stratified sample rather than for a convenience sample.

### *Publication bias*

1  
2 Funnel plots for meta-analysis of CSA and CPA are dominated by between-sample  
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4 heterogeneity, which can both cause asymmetry in the plots or make asymmetry owing to  
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6 publication bias harder to detect [e.g., see 65]. Asymmetry, like any other systematic pattern  
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8 in the data, may also be caused by unaccounted for covariates [66]. For this reason  
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10 asymmetry was assessed for each of the final models using the D-var version of Egger's  
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12 regression test [67]. For neither the CPA final model,  $t(19) = 0.95, p = .35$ , nor the CSA final  
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14 model,  $t(20) = 0.25, p = .81$ , does the regression test detect asymmetry. It is also important to  
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16 consider the mechanism for publication bias, which is likely to be a filtering out of  
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18 statistically non-significant effects via non-submission or biases operating at the publication  
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20 stage. This mechanism is perhaps implausible for CSA and CPA measures as they were  
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22 typically not the primary measures reported in the studies surveyed here and there would be  
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24 no obvious statistical threshold upon which a filter could act.  
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### DISCUSSION

41  
42 This systematic review reveals several striking patterns in the prevalence of childhood  
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44 physical and sexual abuse in adult homeless samples. Most notably, prevalence rates tend to  
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46 be far higher for the homeless samples than for the general population. Thus, this study  
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48 estimated the average prevalence of childhood physical abuse at 37%, which should be  
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50 compared to the estimated rates of physical child abuse to be between 4 and 16% in the  
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52 general populations in the USA, Australia and the UK [26]. Obtaining prevalence assuming  
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54 that a stricter criterion was used (a moderate to extreme cut-off for the CTQ) reduces this  
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1 estimate to 29%.<sup>6</sup> Our review estimated the average prevalence of childhood sexual abuse at  
2 32% for females and 10% for males compared to 7.5% of all children in the general  
3  
4 population (10% for females, 5% for males) [26]. Although there is a huge variation in the  
5  
6 estimates presented in this study, much of the variability is explained through demographic  
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8 factors such as age, gender or ethnicity. For physical, but not sexual abuse, samples with  
9  
10 lower average age tended to have higher prevalence rates. These findings are in line with  
11  
12 previous research that showed childhood physical abuse to be risk factors for homelessness,  
13  
14 with younger homeless people perhaps more likely to have left parental or non-parental care  
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16 to escape abuse [e.g., 46,47,52]. It may also partly explain why one large study with a  
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18 predominately older sample suggested that homeless adults were not disproportionately  
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20 physically abused as children [48]. Also, a small body of research suggests that young people  
21  
22 affected by abuse and neglect risk poor academic achievement at school, which may lead to  
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24 difficulties finding employment in adulthood [26], and, in turn, unemployment is a strong  
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26 cause of homelessness [1-3]. In contrast, older samples likely contain higher proportions of  
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28 homeless caused by other factors (e.g., living conditions associated with poverty) [3,49].  
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30 Furthermore, some variability in measurement and some of the heterogeneity between studies  
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32 is likely to be due to differences in measurement procedures – whether the precise wording of  
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34 items, interpretation of items by participants or rapport with the interviewer.  
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44 Prevalence rates are also higher for more recent studies. While it is tempting to  
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46 conclude that this finding reflects an increase in study quality – with more recent studies  
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48 providing more accurate estimates, it may also be explained by other factors (e.g., an increase  
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<sup>6</sup> Note however that the definitions of CPA and CSA used in general populations research are less strict than implied by the strict criterion implied by a moderate to severe cut-off with the CTQ [see 26, Table 1].

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over time of child abuse in homeless people or reduced stigma leading to greater willingness to disclose; the most recent studies also tended to have small sample sizes).

The moderator analysis suggests a slightly higher prevalence of childhood physical abuse for predominately white samples. Also consistent with previous research, the prevalence of childhood sexual abuse is markedly higher for predominately female samples than for male samples [36, 58]. But while previous research indicated that homeless women who report having been abused as children also report higher rates of violence as adults (see e.g., 59-62), this issue was not investigated in the present study. There is also little or no evidence in our study of an association between race and prevalence of childhood sexual abuse.

Although the moderator analyses present interpretable patterns that are broadly consistent with previous literature, there are potential threats to the validity and generality of the analysis. First, the heterogeneity of the prevalence rates – while not unexpected (given the complex, multifactorial nature of homelessness and childhood abuse) – implies that there remains considerable unexplained between-study variability in prevalence (around 30% for both analyses). Second, the samples are predominately urban and from Western, English-speaking countries and this greatly limits generalizability (given that these are probably atypical homeless populations in a global context). Indeed even within our analyses there was some evidence – albeit weak – that US prevalence rates of sexual abuse among homeless people were slightly lower than for the non-US samples. Taken together with the higher prevalence rates for recent studies, we propose that future research should focus on cross-national, multi-sample studies with common recruitment criteria and common measurement instruments. This may help to eliminate noise in the prevalence estimates between studies and enable other sources of heterogeneity to be measured and modelled.

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Several issues with the retrospective data collection used in the studies included in our analysis need to be considered; the studies may suffer from recall bias because it is impossible to establish whether consequences are due to the actual abuse experience, events that occurred after the abuse experience, or a person's cognitive appraisal of the experience. Other weaknesses of studies using retrospective measurement of childhood abuse include selective inclusion of participants and limited or no opportunity to adjust for social and individual confounding factors as they occur. Our own analyses suggest that convenience samples lead to slightly higher estimates of sexual abuse (but not physical abuse) – potential evidence that sampling strategy does bias estimates. However, it should be pointed out that the majority of homeless research uses retrospective data collection, and recall bias is common to all such studies and that there is no reason to assume that the magnitude or direction of bias is different here. There is further comfort in the broad consistency of estimates between studies that use validated instruments such as the CTQ [42-43,52,57] with those that do not (though – as noted – adopting a strict cut-off on the CTQ does produce lower estimates of CPA).

### **Conclusions**

The substantial heterogeneity between the studies included in the review was not unexpected. One of the findings of this review is that younger people were more likely than older people to report having experienced both physical and sexual abuse during childhood. However, this finding warrants caution as it may be the case that older people report less exposure to trauma than younger people. Longitudinal studies of cohorts of younger and older people who have experienced abuse would help clarify the risk for and routes into homelessness, and identify factors that can prevent homelessness in such populations. Likewise, women were more likely to report childhood sexual abuse compared to men, and

1 longitudinal research can highlight factors that can contribute to homelessness among women  
2 who have been sexually abused as a child. Some of the variation between studies could not be  
3 explained by demographics such as sex, age, or race, and this emphasises the need to  
4 standardize recruitment and measurement procedures (notably measurement instrument) in  
5 order to allow for comparisons across studies as well as study sites.  
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11 This is the first systematic review of childhood physical and sexual abuse in adult  
12 people who are homeless showing that younger individuals and individuals who are white are  
13 more likely to report having experienced child physical abuse compared to those who are  
14 older and non-white and that female homeless are more likely to have experienced sexual  
15 abuse than male homeless regardless of age or ethnicity. As such, the review has important  
16 implications for health services for the homeless. Victims of childhood abuse who are  
17 homeless, or become homeless later in life, have experienced multiple, severe forms of  
18 trauma and the impact on their psychological as well as physical health is often multiple and  
19 severe. This condition is sometimes referred to as *complex posttraumatic disturbance* or  
20 ‘complex trauma’ [68-69]. Complex trauma has been defined as “a combination of early and  
21 late-onset, multiple, and sometimes highly invasive traumatic events, usually of an ongoing,  
22 interpersonal nature” [70, p 1]. Several authors [70-71-73] have recommended trauma  
23 informed care models for people who have experienced complex trauma, including childhood  
24 abuse, and are homeless. These interventions are often based on Herman’s [68] phased  
25 treatment model, where a) the initial phase is the establishment of a therapeutic relationship;  
26 followed by b) improving the safety of trauma sufferers with regard to health, emotions,  
27 relationships, substances and environment; c) helping sufferers to make sense of the  
28 traumatic experience, to understand how the traumatic experience impacts on them in their  
29 current life, and to develop more adaptive ways of coping. From this skills based  
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1 understanding, individuals are then encouraged to make connections within the community  
2 for support in the longer term.  
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4           In conclusion, this study indicates that childhood physical and sexual abuse is more  
5 prevalent among the homeless in Western countries than in the global population. Agencies  
6 who work with people experiencing homelessness should take more account of childhood  
7 trauma among service users.  
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**Conflict of interest**

On behalf of all authors, the corresponding author states that there is no conflict of interest.

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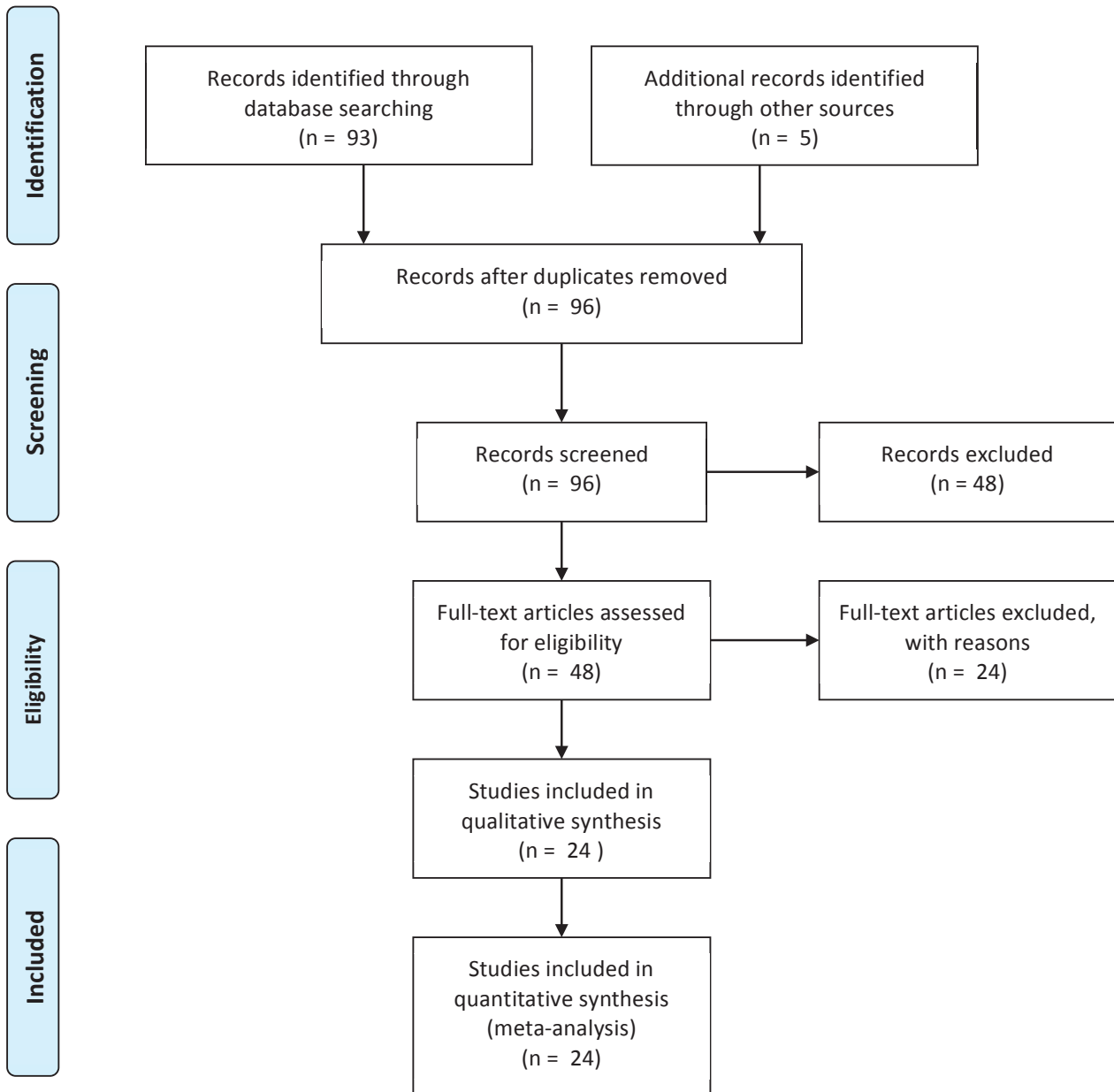
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**Figure 1. Prisma Flow Diagram summarizing selection of studies for the meta-analysis**

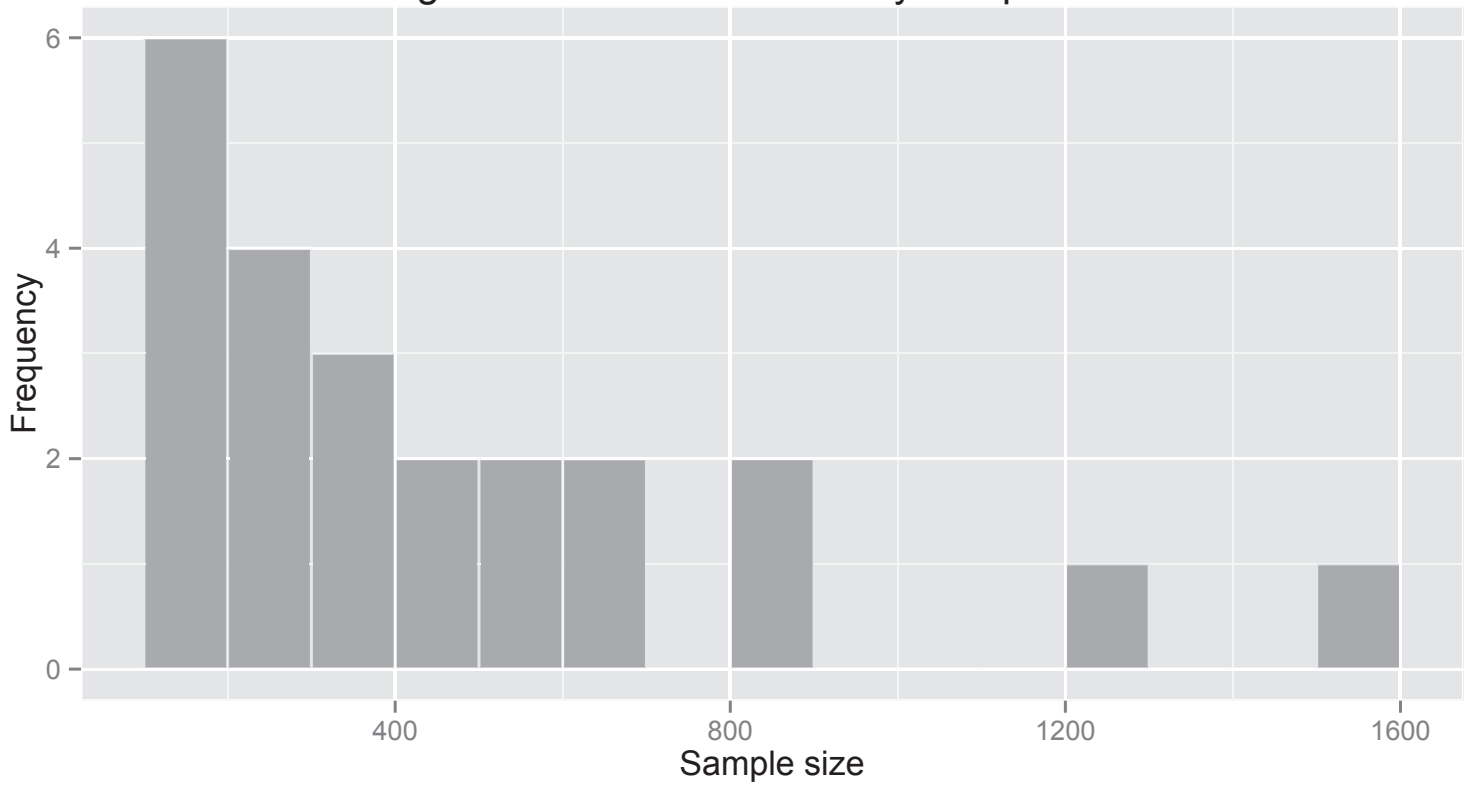


From: Moher D, Liberati A, Tetzlaff J, Altman DG, The PRISMA Group (2009). Preferred Reporting Items for Systematic Reviews and Meta-Analyses: The PRISMA Statement. PLoS Med 6(6): e1000097. doi:10.1371/journal.pmed1000097



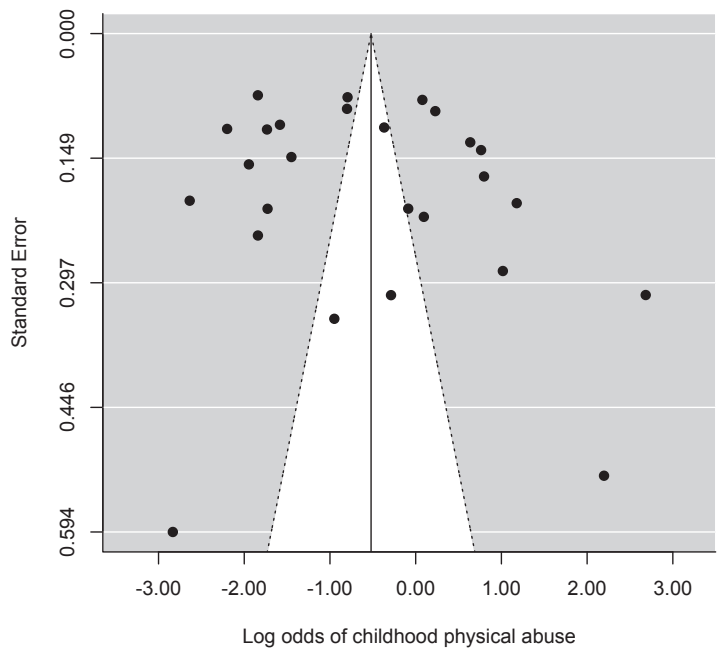
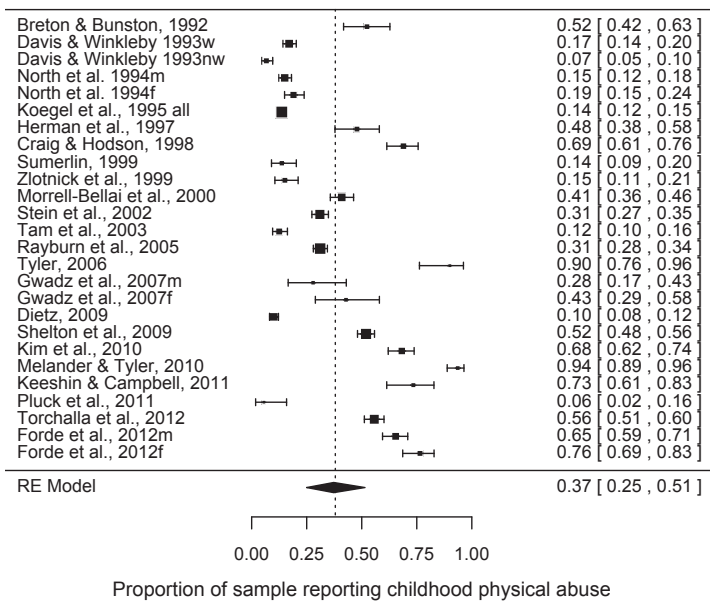
Figure

Figure 2: Distribution of study sample sizes



Figure

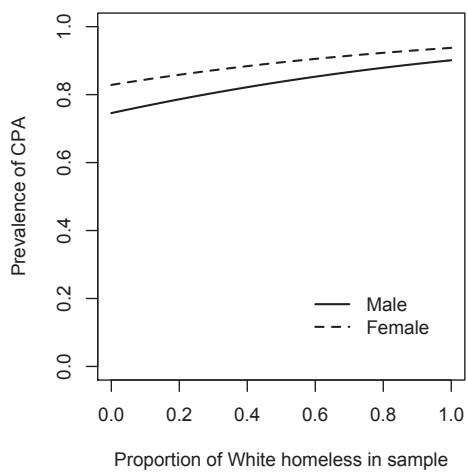
Figure 3. Forest and funnel plots for the meta-analysis of child physical abuse.



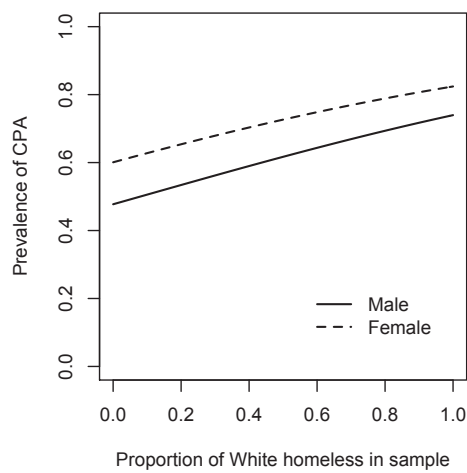
Figure

**Figure 4. Prevalence of physical abuse as a function of demographic factors.**

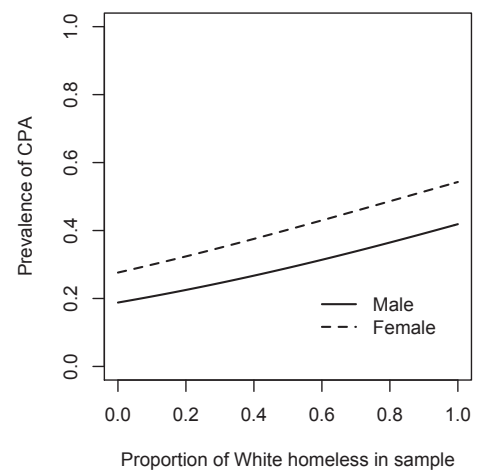
**(a) Average age = 18 years**



**(b) Average age = 35 years**

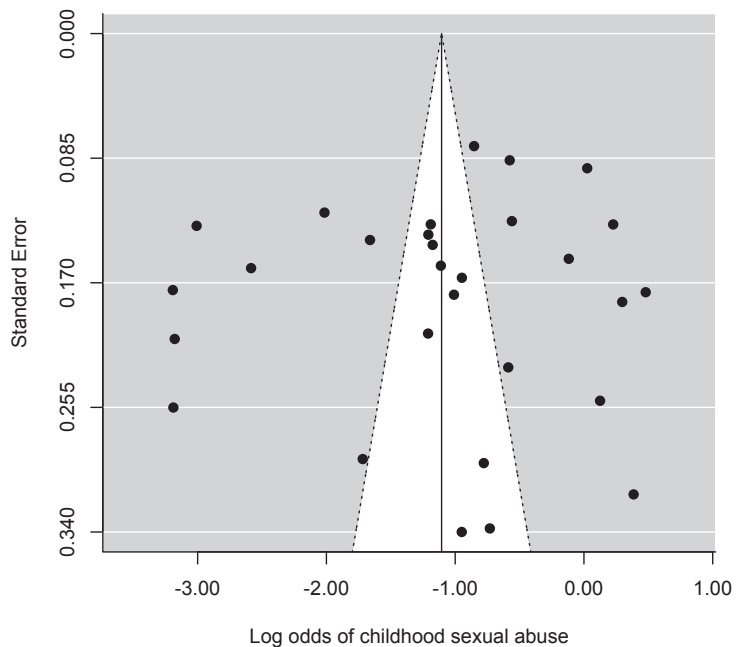
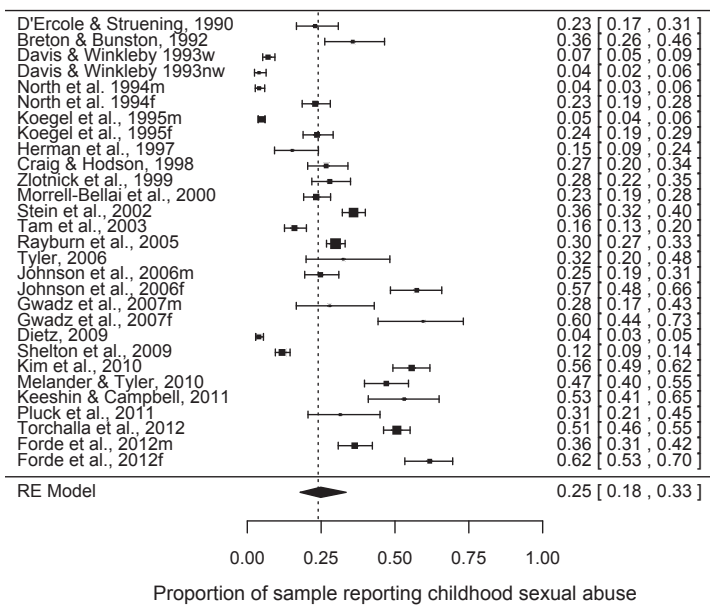


**(c) Average age = 55 years**



Figure

Figure 5. Forest and funnel plots for the meta-analysis of child sexual abuse.



Figure

**Figure 6. Prevalence of sexual abuse as a function of demographic factors.**

