The Development of the Problematic Pornography Consumption Scale (PPCS)

ABSTRACT

Background: To date, no short scale exists with strong psychometric properties that can assess problematic pornography consumption based on an overarching theoretical background. Objectives: The goal of the present study was to develop a brief scale (Problematic Pornography Consumption Scale; PPCS) on the basis of Griffiths’ (2005) six-component addiction model that can distinguish between non-problematic and problematic pornography use. Methods: The PPCS was developed using an online sample of 772 respondents (390 females, 382 males; M_age = 22.56, SD = 4.98 years). Items creation was based on previous problematic pornography use instruments and on the definitions of the factors of Griffiths’ model. Results: A confirmatory factor analysis was carried out—as the scale is based on a well-established theoretical model—leading to an 18-item second-order factor structure. The reliability of the PPCS was excellent and measurement invariance was established. In the current sample, 3.6% of the users belonged to the at-risk group. Based on the sensitivity and specificity analyses we identified an optimal cut-off to distinguish between problematic and non-problematic pornography users. Conclusion: The PPCS is a multidimensional scale of problematic pornography use with strong theoretical basis that also has strong psychometric properties in terms of factor structure and reliability.

Keywords: Behavioral addiction; Pornography use; Problematic Pornography Consumption Scale (PPCS); Psychometric assessment; Scale development
INTRODUCTION

Online pornography consumption is a widespread phenomenon (Edelman, 2009; Haggstrom-Nordin, Hanson, & Tydén, 2005; Hald & Mulya, 2013; Stulhofer, Busko, & Landripet, 2010). Pornography websites are among the top 50 most visited websites worldwide (Alexa.com, 2016; Similarweb.com, 2016) and more than 90% of adults have viewed pornography in their lives (Hald, 2006; Traeen, Spitznogle, & Beverfjord, 2004). In 2016, one of the most popular pornography websites (Pornhub) reported that 4,599,000,000 hours of pornographic videos were watched worldwide. Their statistics also showed that the website was visited approximately 23 billion times meaning that around 44,000 people visited the site every minute (Pornhub.com, 2017). In most of the cases, viewing is not problematic and appears to have little or no negative impact in the person’s life. However, it can become problematic and can have negative effects on one’s life such as problems in romantic relationships or losing one’s job as has been reported in previous studies (e.g., Bergner, & Bridges, 2002; Bostwick & Bucci, 2008; Ford, Durtschi, & Franklin, 2012). In the light of these numbers and findings, it appears to be important to have a multidimensional theory-driven instrument with strong psychometric properties that can assess individual differences in online pornography use in order to distinguish between problematic and non-problematic users and the potential negative consequences of pornography consumption on different groups.

Pornography may mean different things to both researchers and research participants. Therefore, a working definition of pornography is needed prior to assessment (Ayres & Haddock, 2009). However, according to a relatively recent review by Short and colleagues (2012), 84% of the scientific research studies into pornography either did not define pornography and/or did not report whether the research had provided a definition of pornography for their
participants. Hald (2006) used a definition which includes the role of pornography in the creation or enhancement of sexual feelings and thoughts while genitals and/or sexual acts are explicitly shown. This definition was used and refined in later research (Hald & Malamuth, 2008; Reid et al., 2011) and was employed in the present research. According to this definition, “pornography should be defined as material that (i) creates or elicits sexual feelings or thoughts and (ii) contains explicit exposure or descriptions of sexual acts involving the genitals, such as vaginal or anal intercourse, oral sex, or masturbation” (Reid et al., 2011, p. 364).

The assessment of problematic online pornography use is inconsistent indicating that the findings in the area are not comparable (Wéry & Billieux, 2015). According to a recent systematic review (Short et al., 2012), 95% of the researchers used scales and questions that were generated by the authors. Most of the pre-existing psychometric scales did not have a strong theoretical underpinning and they only assessed the frequency of pornography use and/or the time spent using it (e.g., Lam, & Chan, 2007; Lo & Wei, 2005; Meerkerk, Van den Eijnden, & Garretsen, 2006; Stack, Wasserman, & Kern, 2004; Traeen, Spitznogle, & Beverfjord, 2004; Yoder, Virden, & Amin, 2005). In the early 2000s, questionnaires and scales were created that included the topic of problematic online pornography use. However, these mainly concentrated on wider concepts such as sexual addiction, cybersex or the use of internet for sexual purposes (e.g., Carnes & Wilson, 2002; Delmonico & Miller, 2003; Laier, Pawlikowski, Pekal, Schulte, & Brand, 2012). Furthermore, hypersexuality, compulsive pornography use, and compulsive sexual behavior were assessed using several different scales (e.g., Coleman, Miner, Ohlerking, & Raymond, 2001; Noor, Rosser, & Erickson, 2014; Reid, Garos, & Carpenter, 2011; Womack, Hook, Ramos, Davis, & Penberthy; 2013), and only three instruments focused on the narrower concept of problematic pornography use. The nine-item Cyber Pornography Use Inventory
(CPUI-9) was created on the basis of the CPUI-31, therefore the psychometric properties and the factor structure of the CPUI-9 are the only ones taken into consideration here (Wéry & Billieux, 2015; Grubbs et al., 2010, 2015; Kor et al., 2014). The CPUI-9 (Grubbs et al., 2015) has three factors (compulsivity, effort, distress), and the Problematic Pornography Use Scale (PPUS; Kor et al., 2014) has four factors (distress and functional problems, excessive use, control difficulties, and use for escape/avoid negative emotions). Kor et al. (2014) integrated previous problematic pornography, internet use, and hypersexual disorder questionnaires in order to identify these factors. However, as a result of the rather inductive research design, neither the CPUI nor the PPUS has a very strong theoretical background in contrast to other forms of behavioral addiction problematic online behaviors. Furthermore, neither the CPUI nor PPUS included all of the potential dimensions of problematic pornography use (i.e., withdrawal or relapse). The present study aimed to fill this gap by using a deductive strategy and Griffiths’ (2005) components model in order to assess problematic online pornography use because it has been used in the development of many psychometrically robust instruments assessing excessive problematic behavior including social networking (Bánya et al., 2017), gaming (Lemmens, Valkenburg, & Peter, 2009), exercise (Terry et al., 2004), shopping (Andreassen et al., 2015), television series watching (Orosz, Bőthe, & Tóth-Király, 2016), work (Andreassen et al., 2012) and use of Tinder (Orosz, Tóth-Király, Bőthe, & Melher, 2016). Furthermore, neither the CPUI nor PPUS included all of the potential dimensions of problematic pornography use (i.e., withdrawal or relapse). The present study aimed to fill this gap by using a deductive strategy and the six components model of Griffiths (2005) in order to assess problematic online pornography use.

Building on the previous problematic use conceptualizations and scales, the multidimensional Problematic Pornography Consumption Scale (PPCS) was developed on the
theoretical basis of Griffiths’ addiction components model (Griffiths, 2001, 2005). However, it is important to note that the PPCS was established to assess problematic pornography use, not addiction because addiction cannot be assessed on the basis of self-report alone without an in-depth clinical interview (Ross, Mansson, & Daneback, 2012). Accordingly, problematic pornography use included six core elements. The first element is salience referring to high importance of pornography in the person's life that dominates their thinking, feelings and behaviors. The second component refers to mood modification as a subjective experience that users report as a consequence of viewing pornography. This experience can be either arousing or relaxing depending on the desired emotional state. The third dimension is conflict including interpersonal conflicts between problematic users and their significant others, occupational or educational conflicts (depending upon the individual’s age), and intrapsychic conflicts (i.e., knowing the activity is causing problems but feeling they are unable to cut down or cease). The fourth dimension is tolerance and refers to the process whereby increasing amounts of the activity are required to achieve the former mood modifying effects. In the present study, similarly to other arousal behavioral addictions, the quantitative and qualitative aspects of tolerance were focused upon. The quantitative dimension refers to the growing amount of pornography use over time, whereas the qualitative aspect refers to consuming more diverse and extreme pornographic content. According to Zimbardo and Duncan (2012), this qualitative aspect of arousal-based behavioral addictions is related to seeking constantly novel and surprising content. In the case of pornography this can be related to moving from soft-core pornography towards its more extreme hardcore forms. The fifth dimension is related to relapse and is the tendency for repeated reversions to earlier patterns of pornography use and returning to it quickly after abstinence or control. The sixth factor is withdrawal referring to unpleasant
feelings and emotional states that occur when the particular activity is discontinued or suddenly reduced.

As withdrawal and tolerance are usually understood as a consequence of “dependence” (O’Brien, Volkow, & Li, 2006), addiction is a broader construct involving all the six components described above—in line with diagnostic addiction criteria employed in modern psychiatric nosology (American Psychiatric Association, 2013; World Health Organization, 2013). As dependence and addiction are usually viewed as different constructs, the frequency of pornography use and time spent engaging in the activity alone cannot be considered as a satisfactory definition of pornography addiction. It is probable that individuals visit online pornography websites on a very regular basis, but they can stop the activity when it is necessary and that there are few (if any) negative detrimental effects (Kor et al., 2014). Recent research has confirmed this because the relationship between the frequency and duration of pornography use and problematic behavior itself is positive, but only moderate (e.g., Brand et al., 2011; Grubbs, Volk, Exline, & Pargament, 2015; Twohig, Crosby, & Cox, 2009). Addiction and problematic use are overlapping concepts along the same continuum. However, it is more appropriate to use the term “problematic use” instead of “addiction”, when clinical evidence of an actual addiction cannot be provided with the use of self-reported data (Ross et al., 2012).

Considering (i) the pervasive presence of pornography use, (ii) the lack of strongly theory-driven psychometric scale regarding problematic pornography use, and (iii) the lack of potentially important components of problematic pornography use in previous instruments, the goal of the present study was to create a comprehensive psychometric scale that addresses the weakness of previous instruments. Consequently, the aim of the present study was to develop a short, valid, reliable, multidimensional scale that encompasses the most important aspects of
problematic pornography use based on the most extensively tested model of behavioral addictions and problematic online behaviors.

METHODS

Participants and Procedure

The study was conducted in accordance with the Declaration of Helsinki and with the approval of the Institutional Review Board of the research team’s university. The research was conducted via an online questionnaire, and completing it took approximately 15 minutes. Data collection occurred in June 2016 on a public, topic-irrelevant Facebook page that has approximately 217,000 members. Therefore, the collected data were not representative to the population of Hungary. Before starting the questionnaire, participants received detailed information about the study. Subsequently, participants read and approved the informed consent and they also had to indicate that they were 18 years or older.

A total of 1,102 participants were recruited for this research using this online sampling method. Before the analyses, the data was screened and participants were removed for the following reasons: they did not wish to participate in this study (37 individuals), they were under-aged (30 participants), or had the same answer to every questionnaire item (15 individuals). Additionally, those individuals were also excluded who indicated that they had not used pornography in the last six months (248 individuals).

Therefore, a total of 772 participants (females = 390, 50.5%; males = 382, 45.5%) were retained for further analyses and who were aged between 18 and 54 years ($M_{age} = 22.58$, $SD_{age} = 4.89$). Of these participants, 279 lived in a capital city (36.1%), 89 in county towns (11.5%), 286 in towns (37.0%), and 118 in villages (15.3%). Regarding their level of education, 91 had primary school degree (11.8%), 532 had a high school degree (68.9%), and 149 of them had a
degree in higher education (i.e., bachelor, masters or doctoral) (19.3%). Regarding their relationship status, 394 were single (51.0%), 360 were in a relationship (46.6%), and 18 were married (2.3%). Regarding sexual orientation, 621 respondents were heterosexual (80.4%), 82 were heterosexual with homosexuality to some extent (10.6%), 37 were bisexual (4.8%), 10 were homosexual with heterosexuality to some extent (1.3%), 13 were homosexual (1.7%), two were asexual (0.3%), and seven were unsure about their sexual orientation (0.9%). In the last six months, the average frequency of viewing online pornographic videos was weekly and the average time spent with viewing pornography per occasion was 16-30 minutes.

**Measures**

*Problematic Pornography Consumption Scale.* In order to match Griffiths’ (2005) components, the definitions of each component were taken into account. Following this, previous pornography addiction items were considered as potential items in the new instrument (i.e., Grubbs et al., 2010; Kor et al., 2014). However, the strategy of pooling the pre-existing items and analyzing them was not chosen, because the available items (i.e., Grubbs et al., 2010, 2015; Kor et al. 2014) did not include two important components (withdrawal and relapse) and other components were also underrepresented. Finally, in order to have similar wording to other specific and psychometrically robust problematic behavior scales (e.g., Andreassen et al., 2012; Orosz, Bőthe, et al., 2016, Orosz, Tóth-Király, et al., 2016), the items of these scales were considered as a basis of the items of the PPCS. On the basis of these guidelines, a focus group of psychologists (two males and two females, $M_{age} = 27.5$ years, $SD_{age} = 4.65$) created four items per component. To minimize group decision-making biases, an iterative approach was applied. Members first discussed their thoughts in pairs and then in the focus group. Each item had to be:
(a) close to the everyday language used when talking about pornography; (b) easy to understand; (c) concise; (d) clearly belonging to the given dimension, but not to the others; (e) not double-barreled; (h) not suggestive; and (i) adjusted to the scaling. In order to include items that matched Griffiths’ (2005) components, no previous items from alternative problematic pornography instruments remained unchanged. In addition, no previous items in the Griffiths’ model kept the original wording because the subject of the items was replaced with the word “porn” but all other content in the items remained the same. After the focus group created the items, two experts in the addictive behavior field refined the items. In the last step of the item creation, six individuals (young males and females, not psychologists) pretested the items whether they were understandable and close to everyday language use. The final items of the PPCS can be seen in Appendix 1.

Subjective Well-Being Scale (SWB). The SWB scale (Diener, Emmons, Larsen, & Griffin, 1985) is a five-item, one-factor scale assessing the overall satisfaction with life on a 4-point Likert scale, ranging from 1 = “not true to me at all” to 4 = “absolutely true to me” (α = .82).

UCLA Loneliness Scale Version Three. The Revised UCLA Loneliness scale (Russell, 1996) includes 20 items (nine items are reverse-coded) and assesses the feelings of social isolation, lack of connectedness, and the subjective feelings of loneliness (e.g. “How often do you feel that you are no longer close to anyone?”). In the present study, a pretested shortened version of eight items—including reverse-coded items as well—with acceptable validity was used (CFI = .973; TLI = .962; RMSEA = .074 [90% CI = .060-.089]) and reliability (α = .90) (Bőthe, 2016). Respondents rated each item on a four-point scale (1 = never, 4 = always). Higher scores on the scale indicate higher levels of loneliness-related feelings (α = .91).
Relationship satisfaction. Relationship satisfaction was assessed using a single item (“In general, how satisfied are you with your relationship?”) of the Relationship Assessment Scale (Hendrick, 1988; Martos, Sallay, Szabó, Lakatos, & Tóth-Vajna, 2014). Participants had to respond using a 5-point scale (1 = “not satisfied”; 5 = “very satisfied”). This item showed strong positive correlation with the summed score of the Relationship Assessment Scale on previous samples (correlations ranged between .84 and .86), therefore the use of this one-item was deemed sufficient. (For further information, contact the corresponding author).

Sexuality and pornography-related general questions. In addition to standard demographic variables some topic-relevant questions were asked. Sexual satisfaction was asked with one item: “In general how satisfied with your sexual life?” (5-point Likert scale, 1 = “not satisfied”, 5 = “very satisfied”). Frequency of masturbation was asked with one item: “How often do you masturbate?” (9-point Likert scale, 1 = “never”, 9 = “several times a day”). “How often do you watch pornography when you masturbate?” (5-point Likert scale, 1 = “never”, 5 = “very often”). Respondents were asked about their age of the first sexual and pornographic experience. Finally, they were asked about the frequency of reading sexuality-related online stories, viewing pictures, and videos (9-point Likert scale, 1 = “never”, 9 = “several times a day”).

Statistical analysis

For the statistical analysis, SPSS 21 and Mplus 7.3 (Muthén & Muthén, 1998-2012) were used. The initial version of the PPCS comprised 24 items. Each of these items was examined based on three criteria (Fahlman, Mercer-Lynn, Flora, & Eastwood, 2013): (i) corrected item-total correlations, (ii) skewness and kurtosis values for normality, and (iii) content validity compared to other items and the definitions of each problematic use dimension.
After the item selection, confirmatory factor analysis (CFA) was used to assess the dimensionality of the scale. Because the items had severe floor effects in terms of skewness and kurtosis, they were treated as categorical indicators and the mean- and variance-adjusted weighted least squares estimator (WLSMV) was used (Finney & DiStefano, 2006). In the structural assessment, commonly used goodness of fit indices (Brown, 2015; Kline, 2011) were observed with their acceptable or good cut-off values (Bentler, 1990; Brown, 2015; Browne & Cudeck, 1993; Hu & Bentler, 1999; Schermelleh-Engel, Moosbrugger, & Müller, 2003; Tabachnick & Fidell, 2013): the Comparative Fit Index (CFI; ≥ .95 for good, ≥ .90 for acceptable), the Tucker–Lewis index (TLI; ≥ .95 for good, ≥ .90 for acceptable), and the Root-Mean-Square Error of Approximation (RMSEA; ≤ .06 for good, ≤ .08 for acceptable) with a 90% confidence interval. Reliability was assessed using Cronbach’s alpha (Nunnally, 1978).

To test structural invariance between gender groups (male vs. female), several multi-group CFAs were carried out (Meredith, 1993; Vandenberg, 2002; Vandenberg & Lance, 2000). First, the models were estimated freely for both male and female subgroups. Second, four nested models with increasingly constrained parameters were estimated: (1) factor loadings and thresholds were freely estimated (configural invariance), (2) factor loadings were set to be equal (metric invariance), (3) factor loadings and thresholds were set to be equal (scalar invariance), and (4) factor loadings, thresholds, and residual variances were constrained to be equal (residual invariance). Achieving this latter level of invariance is a prerequisite to group-based comparisons based on aggregated manifest scores. When comparing the increasingly constrained models, relative change in fit indices were observed (Chen, 2007; Cheung & Rensvold, 2002; Marsh et al., 2009): ΔCFI ≤ .010; ΔTLI ≤ .010; and ΔRMSEA ≤ .015.
To identify possible groups of pornography users whose activity may be considered problematic, latent profile analysis (LPA) was used and is a person-centered mixture modeling technique that can classify subgroups of people who gave similar responses to the six dimensions (Collins & Lanza, 2010). The analysis was performed with two to four classes on the full sample. To determine the number of latent classes, several indices were used: the Akaike Information Criterion (AIC), the Bayesian Information Criterion (BIC), and the Sample-Size Adjusted Bayesian Information Criterion (SSABIC) where lower values indicate more parsimonious models. Entropy was also examined, indicating the accuracy of the classification process. Higher values indicate higher accuracy with .40 being low, .60 being medium, and .80 being high entropy (Clark & Muthén, 2009). Finally, the Lo-Mendell-Rubin Adjusted Likelihood Ratio Test (L-M-R Test) was also used which compares the estimated model (e.g., three-classes) with a model having one less class (e.g., two-classes). A statistically significant \( p \) value \( (p < .05) \) suggests that the model with more classes fits the data better (Muthén & Muthén, 1998-2012). These groups were then compared along several key variables with analysis of variance (ANOVA) and Bonferroni post-hoc test.

To determine the cut-off point for the PPCS a sensitivity analysis was carried out based on membership in the at-risk group in the LPA. Considering the membership in this group as a “gold standard”, the sensitivity, specificity, positive predictive value, negative predictive value, and accuracy values for all PPCS cut-off points were calculated. Sensitivity was defined as the proportion of true positives belonging to the most problematic group based on the LPA, while specificity was defined as the proportion of the true negatives (Altman & Bland, 1994a; Glaros & Kline, 1988). Positive predictive value was defined as the proportion of the individuals with positive test results that was correctly diagnosed as problematic users, while negative predictive
value was defined as the proportion of participants with negative test results that were correctly
diagnosed as non-problematic users (Glaros & Kline, 1988; Altman & Bland, 1994b).

RESULTS

Dimensionality and structural validity

In the first part of the analysis, each of the initial 24 items were examined based on (a)
their corrected item-total correlations, (b) normality in terms of skewness and kurtosis, and (c)
content validity compared to the other items and pornography viewing in general. Three items
per factor were chosen in order to have a concise and non-repetitive item set. The final items
were chosen as a result of high item-total correlation, relatively low kurtosis, and skewness
values. Furthermore, the aim was to keep the diversity of wording.

Next, CFA was performed on the selected items due to the well-established theoretical
model. The CFA results showed that the theory-based hierarchical model with six factors and a
super-ordinated problematic use dimension (CFI = .977, TLI = .973, RMSEA = .064 [90% CI
.059-.070]) had adequate fit. Factor loadings were high (ranging from .69 to .96) and the six
components loaded strongly on the general factor (ranging from .83 to .92) (see Figure 1). This
18-item six-factor model provides the opportunity to investigate the role of each factor in the
development and maintenance of problematic use.

Figure 1 about here

Measurement invariance
To ensure that group-based comparisons are meaningful, measurement invariance was employed to examine the factor structure of the scale across two subgroups. The results of the invariance analysis are shown in Table 1. In step zero, the baseline models were estimated for both males and females, showing good fit. Then, parameters were gradually constrained and changes in fit indices were observed. In the configural model (M1), all parameters were freely estimated and the fit indices were within the range of acceptability (CFI = .975, TLI = .970, RMSEA = .065 [90% CI .059-.071]). In the metric model (M2), factor loadings were constrained to be equal, resulting in negligible differences in fit indices (ΔCFI = -.002; ΔTLI = .000; ΔRMSEA = .000). In the scalar invariance model (M3), factor loadings and thresholds were set to be equal in both groups, again showing adequacy in terms of fit index changes (ΔCFI = +.001; ΔTLI = +.008; ΔRMSEA = -.009). In the last step, strict invariance model (M4), residual variances were constrained to be equal, and there was no significant deterioration of fit indices compared to the preceding model (ΔCFI = +.003; ΔTLI = +.003; ΔRMSEA = -.005). Fit indices incorporating a control for parsimony (TLI and RMSEA) even resulted in improvements when equality constraints were added, supporting the comparability of the PPCS across gender groups.

Table 1 about here

Gender, age, educational level and place of residence differences

The descriptive statistics of the PPCS are shown in Table 2. PPCS scores weakly correlated with the time spent viewing pornography per occasion (r(770) = .14, p < .01). PPCS correlated with the frequency of reading online pornographic stories (r(770) = .13, p < .01), online pornography picture viewing (r(770) = .27, p < .01), and online pornography video
viewing \((r(770) = .47, p < .01)\). The frequency of masturbation positively correlated with PPCS scores \((r(770) = .38, p < .01)\), and the frequency of pornography consumption during masturbation also positively related with PPCS scores \((r(770) = .27, p < .01)\). The satisfaction with sexual life was weakly and negatively correlated with PPCS scores \((r(372) = -.22, p < .01)\).

One-way ANOVA was used to assess differences in sexual orientation regarding the five larger groups (excluding asexual and unsure respondents as a result of low proportion \(N_{sum} = 8\)). According to the results, no differences in PPCS scores were found regarding sexual orientation. However, gender differences were found as women \((M_{female} = 1.66, SD_{female} = 0.87)\) had lower scores \([t(729.77) = 8.52, p < .01]\) than men \((M_{male} = 2.26, SD_{male} = 1.07)\).

Table 2 about here

*Latent Profile Analysis*

Latent profile analysis was performed on the six PPCS factors. The AIC, BIC, and SSABIC values continuously decreased as more latent classes were added. Regarding entropy, all solutions had high levels of accuracy. The non-significant \(p\) value of the L-M-R Test suggested that the four-class solution should be rejected in favor of the three-class solution (see Table 3). Based on the above-listed criteria, the three-class solution was selected.

Table 3 and Figure 2 about here

The three latent classes with their respective relationship patterns are shown in Figure 2. The *first class* represented non-problematic pornography users (614 individuals, 79.5%). The
second class represented low-risk pornography users (130 individuals, 16.8%). The third class represented at-risk pornography users (28 individuals, 3.6%). The three latent classes and their characteristics can be seen in Table 4.

Table 4 about here

Determination of a potential cut-off score to be classified as a problematic pornography user: sensitivity and specificity analysis

Based on the membership in the third class (i.e., at-risk group) as a “gold standard”, the sensitivity, specificity, positive predictive value (PPV), negative predictive value (NPV), and the accuracy of the PPCS at all possible cut-off points were calculated (Table 5). Based on this analysis, a cut-off score of 76 points was suggested as an optimal cut-off to be classified as problematic pornography user. In this case, sensitivity was 93%, while specificity was 99%. This means that practically 1% of the negative (i.e., non-problematic) cases were considered as problematic, while 7% of the true problematic cases were not recognized. At this value, PPV was 70% and NPV was 100%. This means that 30% of the individuals with a positive test result were identified mistakenly, while all individuals with negative test results were identified correctly. The accuracy of the PPCS was 98%. Increasing the cut-off score would lead to more false negative cases (i.e., problematic pornography users mistakenly diagnosed as non-problematic users), while decreasing the cut-off score would have resulted in more false positive cases (i.e., non-problematic users mistakenly diagnosed as problematic pornography users).
DISCUSSION

The present study aimed to develop a problematic pornography consumption scale that is strongly based on theory alongside robust psychometric properties. Previous scales assessing problematic pornography use either did not have very strong psychometric properties or they had acceptable model fit but the content of the factors raised theoretical questions (Grubbs et al., 2015; Kor et al., 2014). As it is seen from our results, based on theory (Griffiths, 2001; 2005), the PCCS had good factor structure and reliability. This six-factor second-order model provides the opportunity for future research to compare the role of each component in various theoretical frameworks as obsessive vs. harmonious passion towards pornography use (Vallerand, 2015), reward deficiency syndrome (Blum, Cull, Braverman, & Comings, 1996), or motivations regarding pornography use (Reid, Gilliland, Stein, & Fong, 2011).

High levels of invariance (Meredith, 1993; Vandenberg, 2002; Vandenberg & Lance, 2000) were demonstrated across groups formed on the basis of gender (invariance of factor loadings, thresholds and residual variances). The PPCS has strong psychometric properties in terms of factor structure, reliability, and model invariance. According to the latent class analysis, three groups could be reliably distinguished: a non-problematic group, a low-risk group, and an at-risk group. No difference was found regarding sexual orientation, however, similarly to previous studies, males had higher scores on the PPCS than females (Haggstrom-Nordin, Hanson, & Tyden, 2005; Svedin, Akerman, & Priebe, 2011; Traeen et al., 2006).

According to the descriptive statistics, the average participant in the present study viewed pornography-related videos weekly and spent 16-30 minutes viewing pornographic material on each occasion. PPCS scores were weakly related to the time spent viewing pornography, but
moderately to the frequency of viewing pornographic videos. As both time and frequency were asked as categorical variables, it is difficult to calculate a composite score including both. However, the present results suggest that problematic pornography use is more related to the frequency of viewing pornographic videos than the time spent engaging with it on each occasion. Despite the fact that frequent use of pornography is an essential part of problematic pornography use, frequency alone cannot be considered as a satisfactory definition of this phenomenon. It is possible that individuals visit online pornography websites on a regular basis, but they can stop this activity when it is necessary (Kor et al., 2014). Recent research has confirmed this notion, because the relationship between the frequency and duration of use and problematic behavior itself is positive, but only moderate (e.g., Brand et al., 2011; Grubbs et al., 2015; Twohig et al., 2009). Therefore, labeling people as problematic pornography users only on the basis of the duration or the frequency of their pornography consumption is not appropriate.

Furthermore, regarding the form of the pornographic material, the frequency of pornographic video viewing was more strongly related to PPCS scores than viewing pornographic pictures or reading pornographic stories and thus in accordance with previous results (Brand et al., 2011). The frequency of masturbation was also moderately related to problematic pornography use. The strength of this relationship appeared to be even stronger than the association between PPCS scores and the frequency of viewing pornography during masturbation. In line with previous results (e.g., Reid et al., 2011; Reid et al., 2012; Womack et al., 2013), the present results also highlight the relevance of hypersexuality in problematic pornography consumption. More specifically, high level of sexual behavior might be a precursor of problematic pornography use and it is assumed that both problematic pornography use and frequent masturbation are both consequences of hypersexuality. Therefore, problematic
pornography use can appear under the umbrella of hypersexuality similarly to frequent masturbation, going to strip clubs, engaging in telephone sex, and various different forms of cybersex (Kafka, 2010).

Based on latent profile analysis, three severity groups of users were identified. Almost 4% of the sample belonged to the at-risk group. These individuals had high scores on each PPCS component. However, it is important to note that all three groups had relatively lower scores on the conflict component. Arguably, problematic pornography use is not as visible as other forms of problematic behaviors or addictions (such as substance abuse or drinking alcohol). Therefore, the interpersonal conflicts are not as prevalent as in the case of other potentially addictive behaviors. Despite the fact that the at-risk group viewed pornography more frequently and spent more time engaging in it on each occasion, only tendentious differences were found between the low-risk and at-risk groups.

Finally, the sensitivity and specificity analyses revealed an optimal cut-off of 76 points for diagnosing problematic pornography use with the PPCS. However, future studies should further validate this cut-off in a clinical sample in order to consolidate the present findings. Also, it is important to note that the use of scales is limited when employed as an early diagnostic indicator because only clinically-based interview studies are appropriate to diagnose that a specific behavior is truly problematic or pathological for a given individual (Maraz, Király, & Demetrovics, 2015).

The present study is not without limitations. This was a self-selected, self-report questionnaire-based, cross-sectional study that is prone to bias. Furthermore, a longitudinal design would be most useful in examining how potential life events could affect an individual’s problematic pornography use. Although the sample was diverse, and the gender ratio was good,
it was not representative which limits the generalization of the results. Therefore, future studies—similarly to Hald (2006), Luder et al. (2011) and Traeen et al. (2004)—should use representative samples. Regarding the PPCS, the results were based on a correlational design that does not make it possible to infer causality. Further research is needed to examine its temporal stability, as well as convergent, divergent, and predictive validity in different cultures. In terms of clinical practice, prevalence and incidence should be investigated. It would also be useful to examine the relationship patterns between hypersexuality, compulsive behavior and problematic pornography. Further research is also needed to explore whether problematic pornography use and other problematic online behaviors have the same roots. It is possible that these online behaviors have very similar negative consequences.

CONCLUSIONS

The Problematic Pornography Consumption Scale is based on a solid theoretical framework of addictions, specifically using Griffiths’ six-component model (2005), and it has strong psychometric properties in terms of factor structure, reliability, and model invariance. Latent profile analysis identified almost 4% of the sample as at-risk pornography users. However, further clinical investigation and validation are needed to assess the extent of problems related to pornography use. Further cross-cultural research should focus on the characteristics of low- and at-risk groups and identifying potential pathways that lead to problematic pornography use in order to establish potential risk factors and protective factors that can be utilized for prevention and intervention programs.

REFERENCES


or reality? *Archives of Sexual Behavior, 40*(5), 1027-1035. doi: 10.1007/s10508-010-9714-0


Appendix 1.

Problematic Pornography Consumption Scale (PPCS)

Please, think back to the last six months and indicate on the following 7-point scale how often or to what extent the statements apply to you. There is no right or wrong answer. Please indicate the answer that most applies to you.

<table>
<thead>
<tr>
<th>Item</th>
<th>Scores</th>
</tr>
</thead>
<tbody>
<tr>
<td>1. I felt that porn is an important part of my life</td>
<td>0 0 0 0 0 0 0</td>
</tr>
<tr>
<td>2. I used porn to restore the tranquility of my feelings</td>
<td>0 0 0 0 0 0 0</td>
</tr>
<tr>
<td>3. I felt porn caused problems in my sexual life</td>
<td>0 0 0 0 0 0 0</td>
</tr>
<tr>
<td>4. I felt that I had to watch more and more porn for satisfaction</td>
<td>0 0 0 0 0 0 0</td>
</tr>
<tr>
<td>5. I unsuccessfully tried to reduce the amount of porn I watch</td>
<td>0 0 0 0 0 0 0</td>
</tr>
<tr>
<td>6. I became stressed when something prevented me from watching porn</td>
<td>0 0 0 0 0 0 0</td>
</tr>
<tr>
<td>7. I thought about how good it would be to watch porn</td>
<td>0 0 0 0 0 0 0</td>
</tr>
<tr>
<td>8. Watching porn got rid of my negative feelings</td>
<td>0 0 0 0 0 0 0</td>
</tr>
<tr>
<td>9. Watching porn prevented me from bringing out the best in me</td>
<td>0 0 0 0 0 0 0</td>
</tr>
<tr>
<td>10. I felt that I needed more and more porn in order to satisfy my needs</td>
<td>0 0 0 0 0 0 0</td>
</tr>
<tr>
<td>11. When I vowed not to watch porn anymore, I could only do it for a short period of time</td>
<td>0 0 0 0 0 0 0</td>
</tr>
<tr>
<td>12. I became agitated when I was unable to watch porn</td>
<td>0 0 0 0 0 0 0</td>
</tr>
<tr>
<td>13. I continually planned when to watch porn</td>
<td>0 0 0 0 0 0 0</td>
</tr>
<tr>
<td>14. I released my tension by watching porn</td>
<td>0 0 0 0 0 0 0</td>
</tr>
<tr>
<td>15. I neglected other leisure activities as a result of watching porn</td>
<td>0 0 0 0 0 0 0</td>
</tr>
<tr>
<td>16. I gradually watched more “extreme” porn, because the porn I watched before was less satisfying</td>
<td>0 0 0 0 0 0 0</td>
</tr>
<tr>
<td>17. I resisted watching porn for only a little while before I relapsed</td>
<td>0 0 0 0 0 0 0</td>
</tr>
<tr>
<td>18. I missed porn greatly when I didn’t watch it for a while</td>
<td>0 0 0 0 0 0 0</td>
</tr>
</tbody>
</table>

Scoring: Add the scores of the items of each factor. For the total score add all the scores of the items. 76 points or more indicate possible problematic pornography use.

Salience: 1, 7, 13
Mood modification: 2, 8, 14
Conflict: 3, 9, 15
Tolerance: 4, 10, 16
Relapse: 5, 11, 17
Withdrawal: 6, 12, 18