

1 **The occurrence and benefits of post-conflict bystander affiliation in wild Barbary**
2 **macaques (*Macaca sylvanus*)**

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10 Word count = 6083 (excluding references), 2 Figures, 4 Tables.

11 ABSTRACT

12 The majority of studies investigating conflict management in animal societies have
13 focused on the role of reconciliation in mediating the costs of aggression. The function of
14 bystander affiliation (i.e. the selective attraction between an opponent and a bystander in
15 the minutes immediately following aggression) is less well understood.

16 The aim of the current study was to examine, in wild Barbary macaques (*Macaca*
17 *sylvanus*), four potential functions of bystander affiliation with the victim of aggression:
18 1) bystander-initiated affiliation to reduce the victim's post-conflict (PC) anxiety (i.e.
19 'consolation'), 2) victim-initiated affiliation (i.e. 'solicited-consolation'), 3) victim- and
20 bystander-initiated affiliation to avert re-directed aggression (i.e. self-protection), and 4)
21 bystander-initiated affiliation to exploit grooming from the victim. We found partial
22 support for the consolation function as bystander-initiated affiliation occurred more
23 frequently between high quality social partners but had no effect on the victim's PC
24 anxiety. In support of the solicited-consolation function, victim-initiated affiliation
25 occurred more frequently between high quality social partners and also caused a
26 reduction in the victim's PC anxiety. These findings suggest that solicited-consolation
27 may substitute for the stress alleviation role of reconciliation. We found no support for a
28 self-protective function as neither the bystander's or the victim's risk of receiving PC
29 aggression was reduced following bystander affiliation with the victim. Finally,
30 bystanders received significantly more PC grooming than victims, suggesting that
31 grooming exploitation of the victim may drive the bystander's PC behaviour. Our results
32 indicate that bystander affiliation holds different functions and benefits for the victim of

33 aggression and the bystander, and highlights the importance of considering which
34 individual initiates this behaviour.

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36 Key words: Conflict Management; Consolation; Grooming; Reconciliation; Solicited-
37 Consolation; Third-Party Affiliation

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56 In group-living species, conflict between group members is sometimes inevitable as
57 individuals strive for dominance and compete for valuable resources. Opponents
58 experience a number of costs in the minutes immediately following aggression,
59 including, for example, an increased risk of receiving renewed aggression from a former
60 opponent or bystander, elevated post-conflict (PC) anxiety, and reduced feeding
61 opportunities or grooming exchange (Schino 2000; Aureli et al. 2002; McFarland &
62 Majolo 2011b). Therefore, analysing the mechanisms used to mediate the costs of
63 aggression is fundamental to our understanding of how social relationships are
64 maintained in animal societies.

65 Reconciliation, the PC exchange of friendly behaviour between the victim and
66 aggressor (Aureli & de Waal 2000), mediates the costs of aggression by repairing the
67 opponents' social relationship damaged by the conflict, and by reducing their PC anxiety
68 and risk of receiving renewed aggression (Aureli & de Waal 2000). Reconciliation has
69 been demonstrated in over 30 primates (Aureli & de Waal 2000) and several non-primate
70 species (e.g. domestic goats, *Capra hircus*: Schino 2000; wolves, *Canis lupus*: Cordoni &
71 Palagi 2008; ravens, *Corvus corax*: Fraser & Bugnyar 2011). Post-conflict bystander
72 affiliation may also be effective at mediating the opponent's costs of aggression (Fraser
73 et al. 2009). Bystander affiliation is defined as the exchange of friendly behaviour
74 between an opponent and a bystander (i.e. an animal not involved in the former conflict)
75 in the minutes immediately following aggression (Judge 1991). It has been demonstrated
76 in apes (Fraser et al. 2009), wolves (Palagi & Cordoni 2009), ravens (Fraser & Bugnyar
77 2010), rooks (*Corvus frugilegus*: Seed et al. 2007) and horses (*Equus caballus*: Cozzi et
78 al. 2010). However, to date, numerous studies have failed to provide support for the

79 occurrence of bystander affiliation in Old World monkeys (Watts et al. 2000). Moreover,
80 the function of bystander affiliation is less well understood when compared to
81 reconciliation (Aureli et al. 2002; Fraser et al. 2009).

82 Bystander affiliation can be beneficial for the bystander and the victim (Verbeek
83 & de Waal 1997; Fraser et al. 2009). However, studies conducted so far have often failed
84 to take into account the identity of both potential initiators of the affiliation (i.e. victim or
85 bystander). Here we aim to analyse bystander affiliation in wild Barbary macaques
86 (*Macaca sylvanus*) while taking into account the identity of the initiator of the affiliation.
87 Specifically, we aim to test four main, non-mutually exclusive proximate functions of
88 bystander affiliation: 1) Consolation (bystander-initiated), 2) Solicited-consolation
89 (victim-initiated), 3) Self-protection (bystander or victim-initiated), and 4) Exploitation
90 (bystander-initiated). To our knowledge, this is the first study to empirically test these
91 functions of bystander affiliation in a wild non-ape species.

92 Consolation describes the PC scenario whereby bystanders respond to the anxiety
93 of the victim and thus initiate affiliation to appease them (de Waal & Aureli 1996).
94 Despite being a rather anthropomorphic term, consolation might be an innate response in
95 the bystander, elicited by behavioural signs of anxiety in the victim (i.e. self-scratching),
96 that do not involve empathy. We predicted that consolation would reduce PC anxiety in
97 the victim (de Waal & Aureli 1996; Aureli 1997; Wittig & Boesch 2003; Palagi et al.
98 2004). Moreover, we predicted that consolation would occur more frequently between
99 high quality social partners (i.e. between individuals exchanging high rates of affiliation),
100 as bystanders should be more responsive to the signs of anxiety of their friends (Aureli &
101 Schaffner 2002; Fraser et al. 2008a; Fraser & Bugnyar 2010; Romero & de Waal 2010).

102 When testing the solicited-consolation function, we predicted that victims would
103 initiate affiliation with bystanders to reduce their own PC anxiety (de Waal & Aureli
104 1996; Verbeek & de Waal 1997). We also predicted that victims would solicit
105 consolation from bystanders with whom they share high quality relationships as these
106 individuals may be more effective at reducing their PC anxiety (Aureli & Schaffner 2002;
107 Fraser et al. 2008a). Therefore, the consolation and solicited-consolation functions shared
108 similar predictions but differed in the identity of the initiator of PC affiliation (bystander
109 or victim, respectively).

110 For the self-protection functions, we first analysed whether bystanders face an
111 increased risk of receiving re-directed aggression from the victim or aggressor in the PC
112 period (we have previously shown that the victim is at risk of receiving renewed PC
113 aggression from the aggressor or bystander: McFarland & Majolo 2011b). If the
114 bystander or the victim are at risk of receiving PC aggression from each other or from the
115 aggressor (Koski & Sterck 2007), we predicted that the bystanders or the victim,
116 respectively, would affiliate in order to reduce such risk.

117 To our knowledge, the exploitation function has never been tested before (Fraser
118 et al. 2009). For this function, we predicted that more PC grooming would be received by
119 the bystander from the victim, than vice-versa, and that bystanders would target
120 subordinate victims more often than dominants (as subordinate group members tend to
121 give more grooming: Schino 2001; Fruteau et al. 2011). This scenario would be similar to
122 what we found in a previous study (McFarland & Majolo 2011a), showing that, in
123 Barbary macaques, the aggressor often initiates PC affiliation with the victim (i.e.

124 reconciliation) to gain grooming opportunities.

125

126 METHODS

127 *Study subjects and field site*

128 Between September 2008 and August 2009, data were collected daily from 48 individuals
129 living in two groups ('Flat-face' and 'Large') of wild Barbary macaques, in the Middle
130 Atlas Mountains of Morocco (33° 24'N – 005° 12'W). At the beginning of the study,
131 group sizes were 19 (11 males, 8 females) and 29 (19 males, 10 females) adults and sub-
132 adults for the 'Flat-face' and 'Large' group respectively. These groups were non-
133 provisioned and relied on a completely natural diet. Study animals were fully habituated
134 to the presence of researchers (i.e. they did not change their activity when we moved
135 around the study group) and were individually identified via facial characteristics and
136 body size. Permission to conduct our research was granted by the Haut Commissariat des
137 Eaux et Forêts et à la Lutte Contre la Désertification of Morocco. This study complies
138 with Moroccan and UK regulations regarding the ethical treatment of research subjects.

139

140 *Data collection*

141 Data were collected following the post-conflict - matched-control (PC-MC) method (de
142 Waal & Yoshihara 1983; McFarland & Majolo 2011b). The identity and role of the
143 opponents (i.e. aggressor or victim) were recorded anytime aggression was exchanged
144 between two or more individuals. Aggression was recorded anytime at least one of the
145 following behaviours was observed: threat, lunge, chase, slap, grab or bite. The aggressor
146 was defined as the initiator of the first aggressive display. The victim was the recipient of

147 this aggression. Based on the outcome of unidirectional aggressive and submissive
148 interactions collected during baseline focal (see below) and *ad libitum* observations,
149 relative dominance positions (i.e. ranks) were determined for each group member using
150 MatMan 1.0 Software (de Vries et al. 1993). The role of the monkeys in a conflict
151 reflected their dominance relationships, as the aggressor was dominant over the victim in
152 96% of cases (N = 398 of 414 conflicts observed) and only 4% of conflicts involved
153 counter-aggression (i.e. a victim being aggressive towards the former aggressor, N = 17
154 conflicts observed).

155 PC data were collected from either the victim (N = 191) or the aggressor (N = 223)
156 of the conflict for five minutes. PC sessions were postponed if aggression between the
157 former opponents recommenced within 30 seconds of the initial conflict as the conflict
158 was considered to then still be in progress (Aureli 1997). PC data collected from the
159 victim were used to test the bystander-initiated consolation, victim-initiated solicited-
160 consolation, victim- and bystander-initiated self-protection, and the bystander-initiated
161 exploitation functions. PC data collected from the aggressor or the victim were used to
162 test whether bystanders were at risk of receiving PC aggression from the former
163 aggressor or victim of the conflict. During PC sessions we recorded the timing and
164 occurrence of any aggressive or friendly interaction exchanged between the focal
165 opponent and any other group member. We considered grooming, body-contact, mutual
166 teeth-chattering and successful $\leq 1.5\text{m}$ approaches (i.e. approaches that were not followed
167 by aggression or displacement for the first 30 seconds after the approach) as forms of
168 friendly affiliation (Hesler & Fischer 2008; McFarland & Majolo 2011a,b). The initiator
169 (e.g. victim or bystander) of the first PC friendly behaviour was recorded. We recorded

170 all occurrences of self-scratching and used this behaviour as a measure of anxiety. There
171 is comprehensive behavioural, physiological and pharmacological evidence that self-
172 scratching is a reliable measure of anxiety in primates (Schino et al. 1991, 1996;
173 Maestriperi et al. 1992; Barros et al. 2000; Troisi 2002). Moreover, in a previous study
174 on the same study subjects (McFarland & Majolo 2011b) we showed that the victim
175 experienced elevated PC self-scratching rates when compared to MC conditions.

176 MCs were collected within \leq two weeks ($X = 4.63$ days, range = 1 to 14 days) of the
177 matched PCs to control for any variation in the expression of grooming, aggression and
178 self-scratching across the year. To further standardise MC sessions, MCs were only
179 started when, a) the MC focal subject had not been involved in an aggressive interaction
180 with another monkey in the five minutes prior to a planned MC, or during the MC, and,
181 b) no other group member was in close-proximity (i.e. $\leq 1.5\text{m}$) to the MC focal subject.
182 We collected the same data and followed the same methodology during MCs as
183 previously described for the PCs.

184 Scan sampling and focal sampling were used to collect data on the baseline level of
185 affiliation for each dyad. Scan samples were collected every hour on the activity of the
186 study animals (i.e. resting, feeding, allo-grooming, body contact), their $\leq 1.5\text{m}$ proximity
187 to other study subjects, and on the identity of their social partners. Scan data were
188 collected on all subjects visible within ten minutes of the beginning of the scan.
189 Moreover, across the entire study period twenty minute all-occurrences focal sessions
190 were collected from our study animals to determine dyadic dominance relationships, and
191 calculate the proportion of successful $\leq 1.5\text{m}$ approaches exchanged within each dyad.
192 The order of focal sessions on the study animals was randomised each day and focal data

193 were evenly distributed across the study period and time of day. A monkey was never
194 sampled more than once in a single day.

195

196 *Data set and test variables*

197 Of the 414 conflicts analysed, all but one adult male of the “Large group”, and all “Flat-
198 face group” members were targets of at least one PC session ($X = 19$, range = 1 – 31
199 PCs/monkey). 792 scan samples and 1,102 hours of focal observations were collected in
200 the current study ($X = 18.7$, range = 4.7 – 50.9 hours/monkey). Bystander affiliation was
201 defined as the first friendly behaviour (i.e. body-contact, teeth-chattering, grooming)
202 exchanged between the victim and a bystander. Close-proximity approaches were also
203 considered forms of bystander affiliation as there is evidence that close-proximity
204 mediates the costs of aggression in Barbary macaques (McFarland & Majolo, in
205 preparation; Patzelt et al. 2008; McFarland & Majolo 2011a,b). Of the 45 occurrences of
206 bystander affiliation with the victim of aggression, 18 were followed by grooming and 27
207 followed by close-proximity approaches (in the absence of grooming). The occurrence of
208 bystander affiliation was analysed using the ‘PC-MC method’ (de Waal & Yoshihara
209 1983) by comparing the timing of the first friendly behaviour exchanged between the
210 bystander and victim in PC and MC sessions. If a friendly affiliation was not observed
211 during the MC, a conservative latency of 300 seconds was estimated. This estimate was
212 required because if no value was entered for the MC, the PC-MC would have been
213 discarded from the analysis. When friendly behaviour occurred earlier in the PC than the
214 MC (or only in the PC), the PC-MC pair was defined ‘attracted’. When the interaction
215 took place earlier in the MC than in the PC (or only in the MC), the PC-MC pair was

216 defined ‘dispersed’. If the friendly behaviour did not occur in the PC and MC, or if it
217 occurred at the same time, the PC-MC pair was defined ‘neutral’. The proportions of
218 ‘attracted’ and ‘dispersed’ pairs were compared using Wilcoxon signed-ranks tests.

219 When bystander affiliation resulted in grooming, we calculated the percentage of
220 PC grooming received by the victim and bystander. Based on a hypothetical dyad of
221 individual A and B, the percentage of grooming received by individual A (or B) in a
222 grooming bout was calculated using the following equation: [grooming received by A /
223 (grooming received by A + grooming received by B)] x 100. A composite sociality index
224 (CSI) was used to measure the quality of the victim and bystander’s social relationship
225 using the following formula (Silk et al. 2003; McFarland & Majolo 2011b):

$$226 \quad CSI = \frac{\sum_{i=1}^3 \frac{x_i}{m_i}}{3}$$

227 x_i = Dyad’s mean value for each of the three behavioural measures.

228 m_i = Group’s mean value for each of the three behavioural measures.

229

230 Three behavioural variables were entered into this index (exchange of friendly behaviour
231 [i.e. grooming or body-contact], proximity, and tolerance) as they represent three key
232 measures of relationship quality in non-human primates (Fraser et al. 2008b; Majolo et al.
233 2010; Silk et al. 2010; McFarland & Majolo 2011c). To calculate x_i for each dyad we
234 combined data collected from each dyad member on: 1) the proportion of hourly scans
235 each dyad member was exchanging friendly behaviour, 2) the proportion of hourly scans
236 dyad members were within $\leq 1.5\text{m}$ proximity, and, 3) the proportion of successful $\leq 1.5\text{m}$

237 approaches exchanged during the dyad's 20 minute focal sessions. The same three
238 variables were used to calculate medians at the group level to obtain m_i . The higher the
239 CSI value, the stronger the dyad relationship quality was. In this study the values of the
240 CSI ranged from 0 to 8.15 ($X = 1.32$ CSI/dyad).

241

242 *Statistical analysis*

243 We tested our predictions using non-parametric statistics and a series of generalised
244 linear mixed models (GLMMs). To test the consolation, solicited-consolation and self-
245 protection function of bystander affiliation we used three dependent variables in
246 GLMMs: self-scratching, bystander affiliation, and PC aggression received. Two
247 Shapiro-Francia normality tests showed that self-scratching and bystander affiliation
248 were not normally distributed, even after using a square-root transformation. Therefore,
249 these two dependent variables were entered as count data (i.e. N of occurrences in the
250 PC) in GLMMs with Poisson distribution and log link (hereafter Poisson GLMM). In
251 these Poisson GLMMs, the duration in seconds of the PC was the exposure variable in
252 the Poisson GLMM on PC scratching rate. The opportunity to bystander affiliation (i.e.
253 the total number of conflicts involving the victim, excluding those in which the bystander
254 was the opponent of the victim) was as our exposure variable in the Poisson GLMMs on
255 bystander affiliation. For our dichotomous dependent variable (i.e. PC aggression
256 received: yes, no) we used GLMMs with binomial distribution and logit link (hereafter
257 logistic GLMM). Poisson GLMMs do not control for the over-dispersion of the data.
258 Therefore, for each Poisson GLMM we ran a Vuong test (Vuong 1989) to compare the
259 'standard' Poisson GLMM with a zero-inflated Poisson regression. The Vuong tests

260 were all non-significant (see Results below) and thus the results of the Poisson GLMMs
261 were not affected by over-dispersion.

262 GLMMs allow analysing the effect of a series of independent variables (i.e. fixed
263 factors) on a continuous, count or categorical predictor variable (Pinheiro & Bates 2000).
264 Moreover, GLMMs allow analyses to be run using each conflict dyad, or PC and MC
265 session, as a single data point. This procedure is appropriate when using GLMMs, via the
266 inclusion of random factors to the model. Random factors control for the non-
267 independence of the data points (Pinheiro & Bates 2000) thus allowing analyses to be run
268 at the level of the single observation (e.g. PC or MC session) while avoiding any bias due
269 to pseudo-replication.

270 In all the GLMMs presented below, Subject IDs (i.e. victim, aggressor or
271 bystander ID) were entered as ‘crossed’ random factors, thus controlling for pseudo-
272 replication bias at both the individual subject and dyadic level. When comparing PC-MC
273 data, we nested PC-MC pair ID inside Subject ID, as our random factor, so that each PC
274 session was compared to its paired MC. The age combination of the opponents’ dyad
275 (adult-adult, subadult-subadult or adult-subadult), their sex combination (male-male,
276 female-female or male-female), their rank distance, and the occurrence of reconciliation
277 were used as ‘control’ fixed factors because these variables may also play a role in
278 mediating the costs of aggression (Majolo et al. 2009; McFarland & Majolo 2011b).
279 Group ID (‘Flat-face’ or ‘Large’ group) was also entered as a ‘control’ fixed factor. We
280 used this procedure to control for Group ID (instead of entering Group ID as a third
281 random factor with victim and aggressor ID or subject and PC-MC pair ID) because
282 GLMMs could not generate an output for models with three random factors. Note here

283 that entering Group ID as a fixed factor allows testing the effect of an independent
284 variable on a dependent variable while taking into account that the data came from
285 monkeys belonging to different groups (Pineiro & Bates 2000). For a complete list and
286 description of variables used in GLMMS see Table 1. In each GLMM we entered our
287 independent test variable/s together with our control variables. In light of this, we
288 considered the presentation of full GLMM models more comprehensive and conservative
289 to analyse bystander affiliation than the use model selection. All GLMMs were
290 performed in STATA v10.1 software (StataCorp 2007). The exact Wilcoxon tests
291 (Mundy & Fisher 1998) were performed in SPSS Software v17.

292

293 “Approximate location for Table 1”

294

295 *Test models*

296 Model 1: To test whether bystander affiliation reduces PC anxiety in the victim we used
297 data from 191 victim PC sessions. We entered victim PC self-scratching as our dependent
298 variable in a Poisson GLMM where the occurrences of bystander- and victim-initiated
299 bystander affiliation (i.e. yes or no) were the test independent variables (control factors:
300 group ID, dyad age and sex combination, rank difference and the occurrence of
301 reconciliation and bystander-initiated affiliation, random factors: victim and aggressor
302 ID).

303 Models 2 and 3: To test whether bystander affiliation occurs more frequently
304 between high quality social partners we used data based on scores for each group member
305 dyad (N= 450; only dyads with an opportunity of ≥ 1 for bystander affiliation were

306 included in the analysis). We entered either the bystander-initiated affiliation count
307 (Model 2) or the victim-initiated affiliation count (Model 3) as the dependent variable in
308 a Poisson GLMM and dyad relationship quality (i.e. CSI value) as the test independent
309 variable (control factors: group ID, dyad age and sex combination and rank difference,
310 random factor: subject ID,).

311 Models 4 and 5: Based on 200 PC-MC pairs we examined whether bystanders
312 faced an increased PC risk of receiving renewed aggression from the victim or the
313 aggressor compared to MCs. We used two logistic GLMMs on aggression received by
314 bystanders (dichotomous dependent variable, yes or no) from, respectively, the victim
315 (Model 4) or the aggressor (Model 5) and ‘session’ (i.e. PC or MC) as the test
316 independent variable (control factors: group ID, dyad age and sex combination, random
317 factors: PC-MC pair ID nested inside Subject ID).

318 Model 6: To further test the self-protection function we used the 191 PCs
319 collected from the victim and examined whether the occurrence of bystander affiliation
320 reduced re-directed aggression in the bystander or victim, respectively. For the bystander,
321 we ran a logistic GLMM on aggression received by the bystander from the victim (i.e.
322 yes or no) as our dependent variable, and bystander-initiated affiliation (i.e. yes or no) as
323 our test independent variable (control factors: group ID, dyad age and sex combination,
324 rank, reconciliation, victim-initiated affiliation; random factors: victim and aggressor ID).

325 Models 7 and 8: For the victim, we ran two logistic GLMMs on aggression
326 received by the victim from the bystander (Model 7) or the former aggressor (Model 8) as
327 our dependent variable, and victim-initiated affiliation (i.e. yes or no) as our test

328 independent variable (control factors: group ID, dyad age and sex combination, rank,
329 reconciliation, bystander-initiated affiliation; random factors: victim and aggressor ID).

330

331 RESULTS

332 *The occurrence of bystander affiliation*

333 Of the 191 PCs collected from the victim, 24% involved bystander affiliation with the
334 victim; 49% of which were initiated by the bystander, 38% by the victim and 13% were
335 considered to be mutually initiated (i.e. when the bystander and victim approached each
336 other simultaneously). Of the 22 PC-MC pairs involving bystander-initiated affiliation,
337 significantly more pairs were ‘attracted’ (N = 21) compared to those ‘dispersed’ (N = 1)
338 (Wilcoxon: N = 13 subjects, $Z = -2.956$, $P = 0.002$). Of the 17 PC-MC pairs involving
339 victim-initiated affiliation, significantly more pairs were ‘attracted’ (N = 17) compared to
340 those ‘dispersed’ (N = 0) (Wilcoxon: N = 11 subjects, $Z = -3.022$, $P = 0.001$). Therefore,
341 bystander affiliation initiated by the victim or the bystander did occur in Barbary
342 macaques.

343

344 *Consolation*

345 In contrast to our first prediction (i.e. consolation would reduce PC anxiety in the victim),
346 we found no significant difference in the victim’s PC self-scratching following conflicts
347 that resulted in bystander-initiated affiliation or not (Model 1: $\beta \pm SE = -0.159 \pm 0.314$,
348 95% CIs = $-0.774 - 0.456$, $Z = -0.51$, $N = 191$, $P = 0.613$; Vuong test: $z = 0.35$, $P = 0.36$;
349 Fig 1; Table 2).

350

351 “Approximate location for Table 2”

352 “Approximate location for Figure 1”

353

354 In support of the consolation function, bystander-initiated affiliation was more likely to
355 occur in bystander-victim dyads that shared high quality relationships than in those
356 sharing low quality relationships (Model 2: $\beta \pm SE = 0.137 \pm 0.064$, 95% CIs = 0.012 –
357 0.262, $Z = 2.15$, $N = 450$, $P = 0.031$; Vuong test: $z = 1.15$, $P = 0.09$; Table 3). In this
358 analysis, it is important to note that although the count score for bystander affiliation did
359 not control for baseline levels of affiliation for each dyad (Fraser et al. 2008a), there was
360 only one ‘dispersed’ PC-MC pair for bystander affiliation in our dataset. Therefore, it
361 was not considered necessary to adjust these scores according to baseline levels of
362 affiliation as has been done in previous studies (e.g. Fraser et al. 2008a).

363

364 “Approximate location for Table 3”

365

366 *Solicited-consolation*

367 The consolation and solicited-consolation functions shared similar predictions (see
368 above) but differed in being, respectively bystander- or victim-initiated. Therefore, to
369 analyse solicited-consolation we used the same 191 victim PC sessions and 450 group
370 member dyad scores used to test for consolation and similarly structured (in terms of
371 control fixed factors and random factors) Poisson GLMMs as described above.

372 In support of the prediction that victims would initiate affiliation with bystanders
373 to reduce their own PC anxiety, the victim’s PC self-scratching was significantly lower

374 when a conflict was followed by victim-initiated affiliation compared to when not (Model
375 1: $\beta \pm SE = -1.115 \pm 0.519$, 95% CIs = -2.132 – -0.098, $Z = -2.15$, $N = 191$, $P = 0.032$;
376 Vuong test: $z = 0.42$, $P = 0.49$; Fig 1; Table 2). Moreover, victims solicited-consolation
377 more frequently from bystanders with whom they shared high quality relationships
378 (Model 3: $\beta \pm SE = 0.158 \pm 0.074$, 95% CIs = 0.014 – 0.303, $Z = 2.15$, $N = 450$, $P =$
379 0.031; Vuong test: $z = 1.06$, $P = 0.11$; Table 4).

380

381 *Self-protection*

382 We found no significant difference between PC and MCs in the bystander's likelihood of
383 receiving aggression from the victim (Model 4: $\beta \pm SE = 3.052 \pm 1.703$, 95% CIs = -
384 6.390 – 0.286, $Z = -1.79$, $N = 200$, $P = 0.073$; Appendix 1) or the aggressor (Model 5: β
385 $\pm SE = 0.356 \pm 0.491$, 95% CIs = -1.318 – 0.606, $Z = -0.73$, $N = 200$, $P = 0.468$; Appendix
386 2). Therefore, bystanders might not need to affiliate for self-protection as they did not
387 face an increased risk of receiving re-directed aggression from the victim or aggressor.
388 This, however, might still be the case for victims, as victims are at risk of receiving
389 renewed PC aggression from the aggressor or bystander (McFarland & Majolo 2011b).

390 In contrast to the self-protection function, the risk of a bystander receiving re-
391 directed aggression from the victim was not significantly different in the presence or
392 absence of bystander-initiated affiliation (Model 6: $\beta \pm SE = 0.303 \pm 0.808$, 95% CIs = -
393 1.281 – 1.887, $Z = 0.38$, $N = 191$, $P = 0.707$; Appendix 3).

394 The occurrence of victim-initiated affiliation (i.e. yes or no) did not have a
395 significant effect on aggression received by the victim from the bystander (i.e. yes or no)
396 (Model 7: $\beta \pm SE = 0.594 \pm 0.678$, 95% CIs = -0.735 – 1.922, $Z = 0.88$, $N = 191$, $P =$

397 0.381; Appendix 4). Moreover, the victim's risk of receiving renewed aggression from
398 the aggressor was not significantly different in the presence or absence of victim-initiated
399 affiliation (Model 8: $\beta \pm SE = 1.458 \pm 0.875$, 95% CIs = -0.256 – 3.172, $Z = 1.67$, $N =$
400 191, $P = 0.096$; Appendix 5).

401

402 *Exploitation*

403 To test our two predictions for this function (i.e. more PC grooming would be received
404 by the bystander from the victim than vice-versa, and bystanders would target
405 subordinate victims) we used the 17 PCs in which bystander-initiated affiliation was
406 followed by grooming between the victim and the bystander. In support of the
407 exploitation function, bystanders received significantly more grooming than victims in
408 the PC period (Wilcoxon: $N = 18$ victim subjects, $Z = -2.111$, $P = 0.002$). Moreover, out
409 of all the occurrences of bystander affiliation ($N = 45$) we found that bystanders affiliated
410 with subordinate victims significantly more often ($N = 39$, 87%) than they did with
411 dominant victims ($N = 6$, 13%) in the PC period (Wilcoxon: $N = 18$ victim subjects, $Z = -$
412 2.939, $P = 0.002$).

413

414 “Approximate location for Figure 2”

415

416 DISCUSSION

417 Our study is one of a few to have observed bystander affiliation with the victim outside of
418 the great apes, and only the second to have observed this in a macaque species (Watts et
419 al. 2000; Arnold & Barton 2001; Call et al. 2002). In fact, bystander affiliation was a

420 relatively common occurrence in the current study (24% of conflicts involving a focal
421 victim). Through an exploration of four functions of bystander affiliation, we investigated
422 the potential benefits that this PC behaviour offers both the bystander and the victim in
423 the aftermath of a conflict. Unfortunately, kinship data were not available for our study
424 animals and so kin relationships were not considered in our analyses; our results thus
425 have to be interpreted with caution. However, primate social behaviour (e.g. grooming
426 exchange, one of our measures of relationship quality) may be less affected by kinship
427 than originally thought (Schino & Aureli 2010).

428

429 *Why bystanders initiate affiliation with victims of aggression*

430 Consolation is thought to be based on empathy, whereby a bystander initiates contact
431 with a victim in response to the victim's emotional state following aggression (de Waal &
432 van Roosmalen 1979; Fraser et al. 2008a). Acts of consolation are considered to reduce
433 the victim's PC anxiety and to be positively predicted by the quality of the relationship
434 shared by the victim and bystander (Aureli & Schaffner 2002; Fraser et al. 2008a). The
435 current study is the first to directly test the consolation function of bystander affiliation
436 (de Waal & van Roosmalen 1979) in a macaque species. In partial support of this
437 hypothesis, bystanders initiated PC affiliation with victims with whom they shared high
438 quality relationships more frequently than those with low quality relationships. However,
439 we found no evidence that bystander-initiated affiliation reduced the victim's PC anxiety.
440 In the absence of a stress-alleviation effect of bystander affiliation, our findings provide
441 scarce support for the consolation function. This conclusion is in line with the suggestion
442 that non-ape primates do not possess the cognitive capacity for empathy, and thus cannot

443 display consolatory behaviour (de Waal & Aureli 1996). However, bystanders may not
444 necessarily need to be empathic towards the victim's distress in order to affiliate them.
445 An innate response to social or non-social cues (e.g. a conflict or self-scratching) from
446 group companions could elicit bystander affiliation and its potential benefits for the
447 bystander or the victim. For example, the positive link between relationship quality and
448 bystander-initiated affiliation found in this study could result from a tendency for friends
449 (i.e. monkeys sharing a high quality relationship) to maintain proximity while moving,
450 feeding or engaging in other activities. If so, bystanders would be more likely to affiliate
451 victim friends because of their proximity and opportunity to attend to social and non-
452 social cues from the victim, which would elicit a response to such cues without any
453 empathic response.

454 We explored whether bystanders initiate affiliation with the victims of aggression
455 in order to gain grooming opportunities. Bystanders received proportionally more
456 grooming than victims after PC affiliation, similarly to what has previously been found
457 for the aggressor (McFarland & Majolo 2011a). Bystanders also affiliated more
458 frequently with subordinate victims and with victims with whom they shared a high
459 quality relationship. Therefore, exploitation of the victim for grooming appears to be a
460 selective PC tactic whereby bystanders attempt to maximise their grooming return from
461 victims; subordinate monkeys usually give more grooming than they receive (Schino
462 2001; Fruteau et al. 2011) and high quality social partners are generally more 'reliable' or
463 'profitable' grooming partners (Silk et al. 2006, 2010; Schino & Pellegrini 2009).

464 In the PC period when social tension is high, bystanders face an elevated risk of
465 receiving re-directed aggression from the victim. Victims of aggression can re-direct

466 aggression toward bystanders to alleviate stress (Aureli & van Schaik 1991) and deflect
467 the attention of aggression away from themselves (de Waal & van Hooff 1981; Scucchi et
468 al. 1988; Aureli & van Schaik 1991). Therefore, bystanders may affiliate the victim of
469 aggression in order to protect themselves from re-directed aggression (Judge 1991; Aureli
470 & van Schaik 1991; Das 2000; Call et al. 2002; Koski & Sterck 2007). The self-
471 protection function of bystander-initiated affiliation does not explain bystander PC
472 behaviour in wild Barbary macaques, as we found no significant effect of bystander
473 affiliation on aggression received by the bystander.

474

475 *Why victims initiate affiliation with bystanders*

476 Reconciliation is considered to serve a stress alleviating function to the victim whereby
477 exchanging friendly behaviour with their former opponent helps mediate their PC anxiety
478 (Aureli et al. 2002; McFarland & Majolo 2011b). Alternatively, when the risk of receiving
479 renewed aggression from their former opponent is too high, victims may solicit
480 consolation from bystanders as an alternative strategy to mediate their PC anxiety (Watts
481 et al. 2000; Wittig & Boesch 2003). This scenario may apply to our study, as the PC
482 period, even after reconciliation took place, was associated with high rates of renewed
483 inter-opponent aggression (McFarland & Majolo 2011a,b). We found evidence in support
484 of the solicited-consolation function: victim-initiated affiliation reduced their PC anxiety
485 and was predicted by the quality of their relationship with the bystander. Our study is the
486 first to report a stress alleviating function of bystander affiliation in macaques. These
487 novel findings may be due to the fact that we considered the stress alleviating function of
488 bystander affiliation independently for bystander- and victim-initiated affiliation.

489 Whereas in chimpanzees a stress alleviating role of bystander affiliation has been
490 observed in the victim following consolation (i.e. bystander-initiated; Fraser et al. 2008a,
491 but see Koski & Sterck 2007), in Barbary macaques the stress alleviation in the victim is
492 only observed following solicited-consolation (i.e. victim-initiated). Our findings thus
493 evidence the need to consider the identity of the initiator of PC affiliation when exploring
494 the stress alleviation function of bystander affiliation.

495 As for the bystander (see above), we found no evidence for a self-protection
496 function of victim-initiated affiliation. Although victims experienced an increased risk of
497 renewed PC aggression from their former aggressor or bystanders (McFarland & Majolo
498 2011b), victim-initiated affiliation did not reduce such risk. Overall, these findings
499 suggest that bystander affiliation in Barbary macaques does not serve a self-protection
500 function for either the victim or the bystander. Interestingly however, similarly to what
501 has been observed during reconciliation in the same study population (McFarland &
502 Majolo 2011b), solicited-consolation appeared to serve a stress alleviation function (see
503 above) despite the fact that it does not reduce the victim's risk of receiving PC
504 aggression.

505

506 *Conclusions*

507 Our findings highlight the importance of considering whether bystander affiliation is
508 initiated by the victim or the bystander when exploring the function of this PC behaviour.
509 Differences in dominance or resource-holding potential (RHP; Parker 1974) are thought
510 to explain the asymmetric distribution of the costs and benefits of aggression between
511 victims and aggressors (e.g. Schino et al. 2007; Cooper et al. 2007; Koski et al. 2007;

512 Schino et al. 2007; McFarland & Majolo 2011b) as well as their PC social tactics
513 (McFarland & Majolo 2011a). Similar asymmetries are expected to occur between the
514 bystander and the victim. Therefore, the decision-making processes made by the victim
515 or bystander to affiliate following a conflict, are potentially driven by different ‘motives’
516 and benefits: victims attempt to reduce their PC anxiety whereas bystanders benefit from
517 grooming opportunities. Both benefits are more likely to be gained once the victim or the
518 bystander initiate PC affiliation with a high quality social partner. Although the
519 importance of considering the initiator of bystander affiliation has long been recognised
520 (de Waal & Aureli 1996; Verbeek & de Waal 1997; Fraser et al. 2008a), the majority of
521 previous studies have failed to account for this important parameter in studies of conflict
522 management. We propose that when testing for the occurrence bystander affiliation,
523 identifying the initiator of these interactions is crucial to further understanding of its
524 functional significance, as well as make sure the correct functional hypothesis is being
525 tested. Moreover, the lack of distinction between bystander-initiated and victim-initiated
526 affiliation in previous studies may explain the lack of evidence for the different functions
527 of bystander affiliation in animal societies (Fraser et al. 2009).

528 Bystander affiliation has been described as a mutualistic behaviour whereby
529 benefits are reciprocated between the victim and bystander (Aureli et al. in press). Our
530 findings support this view as bystander affiliation provides a stress-alleviation benefit to
531 the victim, and grooming benefits to the bystander. The adaptive value of bystander
532 affiliation appears to be two-fold. Bystander affiliation is used by the victim or bystander
533 to manage the costs of aggression and to maintain the benefits of high quality social
534 relationships, both of which impact on an individual’s physiological well-being and

535 fitness (Keverne et al. 1989; van Schaik & Aureli 2000; Silk et al. 2003, 2009, 2010).
536 Bystander affiliation also appears to be used by bystanders as a means to receive
537 grooming and its social and hygienic benefits (Keverne et al. 1989; Zamma 2002; Dunbar
538 2010).

539

540 ACKNOWLEDGEMENTS

541 We are extremely grateful to Chris Young, Laëtitia Maréchal, Pawel Fedurek and Paolo
542 Piedimonte for their invaluable assistance in the field. We would also like to thank
543 Professor Mohammed Qarro (Ecole Nationale Forestière d'Ingénieurs, Morocco) for his
544 support in the field. We thank Sami Merilaita, Gabriele Schino, Daniel Stahl and three
545 anonymous reviewers for useful comments on previous versions of this manuscript.

546

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711 FIGURES

712 Figure 1. Box-plot (median, range, upper and lower quartiles) showing the victim's post-
713 conflict self-scratching count in the presence or absence of bystander- or victim-initiated
714 affiliation

715

716 Figure 2. Box-plot (median, range, upper and lower quartiles) showing the percentage of
717 post-conflict grooming received by victims and bystanders

718

719 TABLES

720 Table 1. Variables used in the GLMMs (see Methods and Results for details on which

721 variables were used to test each prediction of this study).

| Name | Type |
|--|--|
| <i>Dependent variables</i> | |
| PC self-scratching | Count |
| Bystander affiliation tendency (initiated by the victim or bystander) | Count |
| PC aggression received | Binomial (i.e. yes or no) |
| <i>Independent variables</i> | |
| Bystander-initiated affiliation | Binomial (i.e. yes or no) |
| Victim-initiated affiliation | Binomial (i.e. yes or no) |
| Composite sociality index | Continuous |
| PC-MC session | Binomial (i.e. PC or MC session) |
| <i>Control variables</i> | |
| Group | Binomial (i.e. 'Flat-face' or 'Large' group) |
| Age combination | Binomial (i.e. adult-adult or adult-subadult dyads) |
| Sex combination | Binomial (i.e. same sexed or different sexed dyads) |
| Rank difference | Continuous |
| Reconciliation | Binomial (i.e. yes or no) |
| <i>Random factors</i> | |

| | |
|--------------|--|
| Victim ID | Multinomial (ID number of the conflict victim) |
| Aggressor ID | Multinomial (ID number of the conflict aggressor) |
| Subject IDs | Multinomial (ID number of individuals in each group member dyad) |
| PC-MC pair | Multinomial (ID number of each PC-MC pair) |

723 Table 2. GLMM Poisson-regression results for the relationship between victim post-
 724 conflict self-scratching count and bystander affiliation (initiated by the bystander or
 725 victim; N = 191) [Model 1]

726

| | $\beta \pm SE$ | <i>Z</i> | <i>P</i> | 95% CIs |
|--|------------------|----------|----------|-------------------|
| Group | -0.2884 ± 0.2391 | -1.21 | 0.228 | -0.7570 - 0.1801 |
| Age combination | 0.1174 ± 0.1561 | 0.75 | 0.452 | -0.1885 - 0.4234 |
| Sex combination | 0.0152 ± 0.2205 | 0.07 | 0.945 | -0.4170 - 0.4473 |
| Rank difference | 0.0271 ± 0.0198 | 1.37 | 0.172 | -0.0118 - 0.0659 |
| Reconciliation | -0.7366 ± 0.3284 | -2.24 | 0.025 | -1.3802 - -0.0930 |
| Bystander-initiated affiliation | -0.1588 ± 0.3140 | -0.51 | 0.613 | -0.7742 - 0.4565 |
| Victim-initiated affiliation | -1.1147 ± 0.5188 | -2.15 | 0.032 | -2.1315 - -0.0979 |
| <i>Random effects</i> | | | | |
| Victim ID estimated variance ± SE = 0.1436 ± 0.3220 | | | | |
| Aggressor ID estimate variance ± SE = 0.00001 ± 0.2218 | | | | |

727

728 Table 3. GLMM Poisson-regression results for the relationship between bystander-
 729 initiated affiliation count and bystander-victim relationship quality (N = 450) [Model 2]

730

| | $\beta \pm SE$ | <i>Z</i> | <i>P</i> | 95% CIs |
|---------------------------|----------------------|----------|----------|------------------|
| Group | 0.4410 \pm 0.5957 | 0.74 | 0.459 | -0.7266 - 1.6086 |
| Age combination | -0.9262 \pm 0.7530 | -1.23 | 0.219 | -2.4021 - 0.5497 |
| Sex combination | 0.1228 \pm 0.5082 | 0.24 | 0.809 | -0.8733 - 1.1188 |
| Rank difference | 0.04707 \pm 0.0382 | 1.23 | 0.218 | -0.0279 - 0.1220 |
| Dyad relationship quality | 0.1369 \pm 0.0636 | 2.15 | 0.031 | 0.0122 - 0.2616 |

Random effects

Victim ID estimated variance \pm SE = 0.0947 \pm 0.3283

Aggressor ID estimated variance \pm SE = 0.3231 \pm 0.1781

731
732

733 Table 4. GLMM Poisson-regression results for the relationship between victim-initiated
 734 affiliation count and bystander-victim relationship quality (N = 450) [Model 3]

735

| | $\beta \pm SE$ | <i>Z</i> | <i>P</i> | 95% CIs |
|--|----------------------|----------|----------|------------------|
| Group | 0.0697 \pm 0.6982 | 0.1 | 0.921 | -1.2988 - 1.4381 |
| Age combination | -0.4312 \pm 0.7374 | -0.58 | 0.559 | -1.8765 - 1.0141 |
| Sex combination | -0.3592 \pm 0.5360 | -0.67 | 0.503 | -1.4098 - 0.6914 |
| Rank difference | 0.0809 \pm 0.0437 | 1.85 | 0.064 | -0.0046 - 0.1665 |
| Dyad relationship quality | 0.1584 \pm 0.0737 | 2.15 | 0.031 | 0.0141 - 0.3026 |
| <i>Random effects</i> | | | | |
| Subject A ID estimated variance \pm SE = 0.4248 \pm 0.4222 | | | | |
| Subject B ID estimated variance \pm SE = 0.4357 \pm 0.4455 | | | | |

736

737

738 APPENDICES

739 Appendix 1. GLMM logistic-regression results for the relationship between aggression

740 received by bystanders from victims and PC-MC session (N = 200) [Model 4]

741

| | $\beta \pm SE$ | <i>Z</i> | <i>P</i> | 95% CIs |
|---|----------------------|----------|----------|-------------------|
| Group | -2.6275 \pm 3.2899 | -0.8 | 0.424 | -9.0757 - 3.8206 |
| Age combination | -3.9583 \pm 4.0410 | -0.98 | 0.327 | -11.8785 - 3.9619 |
| Sex combination | -3.2627 \pm 3.2304 | -1.01 | 0.312 | -9.5941 - 3.0687 |
| PC-MC session | -3.0519 \pm 1.7030 | -1.79 | 0.073 | -6.3897 - 0.2858 |
| <i>Random effects</i> | | | | |
| Subject ID (nested PC-MC pair ID) estimated variance \pm SE = 8.1534 \pm 3.5931 | | | | |

742

743 Appendix 2. GLMM logistic-regression results for the relationship between aggression
 744 received by bystanders from aggressors and PC-MC session (N = 200) [Model 5]

745

| | $\beta \pm SE$ | <i>Z</i> | <i>P</i> | 95% CIs |
|---|------------------|----------|----------|------------------|
| Group | 0.2864 ± 0.5038 | 0.57 | 0.57 | -0.7009 - 1.2737 |
| Age combination | 0.5793 ± 1.1652 | 0.5 | 0.619 | -1.7044 - 2.8630 |
| Sex combination | -0.7432 ± 0.5947 | -1.25 | 0.211 | -1.9088 - 0.4224 |
| PC-MC session | -0.3560 ± 0.4910 | -0.73 | 0.468 | -1.3182 - 0.606 |
| <i>Random effects</i> | | | | |
| Subject ID (nested PC-MC pair ID) estimated variance ± SE = 0.0009 ± 0.6976 | | | | |

746

747 Appendix 3. GLMM logistic-regression results for the relationship between aggression
 748 received by bystanders from victims and bystander-initiated affiliation (N = 191) [Model
 749 6]

750

| | $\beta \pm SE$ | <i>Z</i> | <i>P</i> | 95% CIs |
|---------------------------------|------------------|----------|----------|-------------------|
| Group | -0.8125 ± 0.7987 | -1.02 | 0.309 | -2.3780 - 0.7529 |
| Age combination | -1.2780 ± 0.5999 | -2.13 | 0.033 | -2.4537 - -0.1022 |
| Sex combination | 0.2969 ± 0.6925 | 0.43 | 0.668 | -1.0604 - 1.6541 |
| Rank difference | 0.09545 ± 0.0573 | 1.66 | 0.096 | -0.0169 - 0.2079 |
| Reconciliation | 0.2346 ± 0.7280 | 0.32 | 0.747 | -1.1922 - 1.6615 |
| Bystander-initiated affiliation | 0.3034 ± 0.8081 | 0.38 | 0.707 | -1.2805 - 1.8873 |
| Victim-initiated affiliation | 1.4173 ± 0.7695 | 1.84 | 0.066 | -0.0909 - 2.9254 |

Random effects

Victim ID estimated variance ± SE = 0.9094 ± 0.4376

Aggressor ID estimated variance ± SE = 0.4868 ± 0.6295

751

752 Appendix 4. GLMM logistic-regression results for the relationship between aggression
 753 received by victims from bystanders and victim-initiated affiliation (N = 191) [Model 7]

754

| | $\beta \pm SE$ | <i>Z</i> | <i>P</i> | 95% CIs |
|---|------------------|----------|----------|------------------|
| Group | -1.0079 ± 0.6328 | -1.59 | 0.111 | -2.2481 - 0.2323 |
| Age combination | -0.0508 ± 0.3821 | -0.13 | 0.894 | -0.7996 - 0.6980 |
| Sex combination | -0.0937 ± 0.5334 | -0.18 | 0.861 | -1.1393 - 0.9518 |
| Rank difference | 0.04736 ± 0.0525 | 0.9 | 0.367 | -0.0555 - 0.1502 |
| Reconciliation | -0.3903 ± 0.7031 | -0.56 | 0.579 | -1.7683 - 0.9877 |
| Bystander-initiated affiliation | -0.3546 ± 0.8258 | -0.43 | 0.668 | -1.9732 - 1.2640 |
| Victim-initiated affiliation | 0.5939 ± 0.6779 | 0.88 | 0.381 | -0.7347 - 1.922 |
| <i>Random effects</i> | | | | |
| Victim ID estimated variance ± SE = 3.03e-09 ± 0.4033 | | | | |
| Aggressor ID estimated variance ± SE = 0.6383581 ± 0.4225 | | | | |

755

756 Appendix 5. GLMM logistic-regression results for the relationship between aggression
 757 received by victims from aggressors and victim-initiated affiliation (N = 191) [Model 8]

758

| | $\beta \pm SE$ | <i>Z</i> | <i>P</i> | 95% CIs |
|---------------------------------|------------------|----------|----------|-------------------|
| Group | 0.0412 ± 0.6698 | 0.06 | 0.951 | -1.2715 - 1.3539 |
| Age combination | -0.1594 ± 0.4364 | -0.37 | 0.715 | -1.0147 - 0.6960 |
| Sex combination | -1.5156 ± 0.7389 | -2.05 | 0.04 | -2.9639 - -0.0673 |
| Rank difference | 0.0413 ± 0.059 | 0.7 | 0.483 | -0.0741 - 0.1577 |
| Reconciliation | 1.2469 ± 0.6796 | 1.83 | 0.067 | -0.0850 - 2.5788 |
| Bystander-initiated affiliation | 1.2027 ± 0.7943 | 1.51 | 0.13 | -0.3542 - 2.7596 |
| Victim-initiated affiliation | 1.4580 ± 0.8746 | 1.67 | 0.096 | -0.2562 - 3.1722 |

Random effects

Victim ID estimated variance ± SE = 0.4836 ± 0.7469

Aggressor ID estimated variance ± SE = 1.54e-06 ± 1.1251

759